Asymptotics of the Sketched Pseudoinverse

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Age=27, Height=5'11", ABO=A, ALT=36, Glucose=78, Creatinine=0.99, Sodium=132, Carbon Dioxide=21...

Age=56, Height=5'3", ABO=AB, ALT=40, Glucose=98, Creatinine=0.63, Sodium=182, Carbon Dioxide=25...

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- Random forests [1]
 - Ensemble of decision trees
 - Trained on bootstrap samples
 - Branch on random feature subsets



Existing Practice

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 - Ensemble of decision trees
 - Trained on **bootstrap samples**
 - Branch on random feature subsets
- Random projection ensembles [2]
 - Ensemble of sketched regressors
 - Randomly project observations
 - Randomly project features



Existing Practice

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 - Random forests are difficult to analyze
 - E.g., purely random forests [3] make analysis tractable
 - Simplified models [e.g., 2, 3] have proven upper bounds
 - Consistent estimation
 - Ensembles strictly better than individual models

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- If there is implicit regularization in ensembles, can we make it explicit?
- Stronger than consistency, tighter than upper bounds
 - Understand both good/optimal as well as suboptimal ensemble models

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Spoiler:

Randomized least squares = ridge + noise

Randomized ensembles = ridge

• New question: Wha

- If there is implicit regularization in ensembles, can we make it explicit?
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- Ground truth function $f\colon \mathcal{X} o \mathbb{R}$
- Ensemble of estimators $\hat{f}_1, \ldots, \hat{f}_K \colon \mathcal{X} \to \mathbb{R}$
- Squared error

$$R(\hat{f}_1, \dots, \hat{f}_K) \triangleq \mathbb{E} \left[R(\hat{f}_1, \dots, \hat{f}_K; x) \right]$$
$$R(\hat{f}_1, \dots, \hat{f}_K; x) \triangleq \mathbb{E} \left[\left(\frac{1}{K} \sum_{k=1}^K \hat{f}_k(x) - f(x) \right)^2 \right]$$

$$\mathbf{y} = (y_1, \dots, y_n)^ op \in \mathbb{R}^n$$

Early Work: Linear Regression Setting

[4] DL, H Javadi, RG Baraniuk. "The implicit regularization of ordinary least squares ensembles." AISTATS, 2020.

• Training data: $\mathbf{X} = \begin{bmatrix} \mathbf{x}_1 \dots \mathbf{x}_n \end{bmatrix}^\top \in \mathbb{R}^{n \times p}$

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• Data model:
$$\mathbf{x}_i \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(\mathbf{0}, \mathbf{I}_p), \ f(\mathbf{x}) = \boldsymbol{\beta}^{*\top} \mathbf{x}, \ \|\boldsymbol{\beta}^*\|_2 = 1$$

 $y_i \sim \mathcal{N}(f(\mathbf{x}_i), \sigma^2)$

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• Estimators:
$$\hat{f}_k(\mathbf{x}) = \widehat{\boldsymbol{\beta}}_k^\top \mathbf{x}$$

 $\widehat{\boldsymbol{\beta}}_k = \mathbf{S}_k \cdot \operatorname*{arg\,min}_{\mathbf{b}} \frac{1}{2} \| \mathbf{T}_k^\top (\mathbf{y} - \mathbf{X}\mathbf{S}_k \mathbf{b}) \|_2^2$

Random Subsampling

- Subsampling operators $\mathbf{S}_k \in \mathbb{R}^{p \times q}, \ \mathbf{T}_k \in \mathbb{R}^{n \times m}, \ m > q + 1$
 - Uniformly random columns of identity matrix



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Turn the Crank...





Error under Proportional Asymptotics

• Exact error expression for finite ensembles

Theorem 1. In the limit as $(n, p, m, q) \to \infty$ with $p/n \to \gamma$, $m/n \to \eta$, $q/p \to \alpha$, if $\eta > \alpha \gamma$,

$$R(\hat{f}_1,\ldots,\hat{f}_K) = \frac{K-1}{K} \left(\frac{(1-\alpha)^2 + \sigma^2 \alpha^2 \gamma}{1-\alpha^2 \gamma} \right) + \frac{1}{K} \left(\frac{\eta(1-\alpha) + \sigma^2 \alpha \gamma}{\eta - \alpha \gamma} \right).$$

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Infinite ensemble error depends only on feature subsampling

$$R_{\mathrm{ens}}^{\infty}(\boldsymbol{\alpha}) \triangleq \lim_{K \to \infty} R(\hat{f}_1, \dots, \hat{f}_K) = \frac{(1-\boldsymbol{\alpha})^2 + \sigma^2 \boldsymbol{\alpha}^2 \gamma}{1-\boldsymbol{\alpha}^2 \gamma}$$

• Ridge regression: $\hat{f}_{\lambda}(\mathbf{x}) = \widehat{\boldsymbol{\beta}}_{\lambda}^{\top}\mathbf{x}, \ \widehat{\boldsymbol{\beta}}_{\lambda} = \operatorname*{arg\,min}_{\mathbf{b}} \frac{1}{2n} \|\mathbf{y} - \mathbf{X}\mathbf{b}\|_{2}^{2} + \frac{\lambda}{2} \|\mathbf{b}\|_{2}^{2}$

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- Tuning feature subsampling is optimal

Theorem 2. Under the conditions of Theorem 1 and if $\beta^* \sim \mathcal{N}(\mathbf{0}, p^{-1}\mathbf{I})$,

$$\inf_{\alpha < \gamma^{-1}} R^{\infty}_{\text{ens}}(\alpha) = \inf_{\lambda > 0} R(\hat{f}_{\lambda})$$

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• However, proof sheds little insight

$$\inf_{\alpha < \gamma^{-1}} R^{\infty}_{\text{ens}}(\alpha) = \frac{1}{2} \left(\frac{\gamma - 1}{\gamma} - \sigma^2 + \sqrt{\left(\sigma^2 - \frac{\gamma - 1}{\gamma}\right)^2 + 4\sigma^2} \right) = \inf_{\lambda > 0} R(\hat{f}_{\lambda})$$

Interpretation of Results

- Since ridge regression is optimal, ensemble is optimal
 - Ridge regression is the minimum mean squared error (MMSE) estimator

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- Does this imply ensemble converges to ridge regression?
 - For finite dimensions, MMSE is unique
 - For infinite dimensions?
 - Optimality theorem suggests convergence to ridge, but not rigorous

Interpretation of Results

- Since ridge regression is optimal, ensemble is optimal
 - Ridge regression is the minimum mean squared error (MMSE) estimator
- Does this imply ensemble converges to ridge regression?
 - For finite dimensions, MM
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 - Optimality theorem su

But what is this predictor?

Jut not rigorous



• Infinite ensemble with only feature subsampling:

$$\hat{f}_{\text{ens}}^{\infty}(\mathbf{x}) = \lim_{K \to \infty} \frac{1}{K} \sum_{k=1}^{K} \widehat{\boldsymbol{\beta}}_{k}^{\top} \mathbf{x} = \mathbf{y}^{\top} \mathbf{X} \left(\lim_{K \to \infty} \frac{1}{K} \sum_{k=1}^{K} \mathbf{S}_{k} \left(\mathbf{S}_{k}^{\top} \mathbf{X}^{\top} \mathbf{X} \mathbf{S}_{k} \right)^{-1} \mathbf{S}_{k}^{\top} \right) \mathbf{x}$$

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Determinantal Point Process pseudoinverse [7]:

Theorem (Mutny et al. 2020). If $\mathbf{M} \succ \mathbf{0}$ and $\mathbf{S} \sim \mathcal{DPP}(\frac{1}{\lambda}\mathbf{M})$,

$$\mathbb{E}\left[\mathbf{S}\left(\mathbf{S}^{ op}\mathbf{M}\mathbf{S}
ight)^{-1}\mathbf{S}^{ op}
ight]=\left(\mathbf{M}+\lambda\mathbf{I}
ight)^{-1},$$

where λ is the solution to $\mathbb{E}\left[\frac{q_{\mathbf{S}}}{p}\right] = \operatorname{tr}(\mathbf{M}\left(\mathbf{M} + \lambda \mathbf{I}\right)^{-1}).$

[5] M Mutny, M Dereziński, A Krause. "Convergence analysis of block coordinate algorithms with determinantal sampling." AISTATS, 2020.

Moving to Sketched Ensembles

- Problem: strict isotropic assumption on data
 - Solution: let subsampling operators be isotropic sketches instead

$$[\mathbf{S}]_{ij} \overset{\text{i.i.d.}}{\sim} \mathcal{N}(0, \frac{1}{q}), \ [\mathbf{T}]_{ij} \overset{\text{i.i.d.}}{\sim} \mathcal{N}(0, \frac{1}{m})$$

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- Problem: estimators use ordinary least squares, m > q + 1
 - Solution: consider ridge regression for arbitrary sketch sizes

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• Asymptotic equivalences [6]:

Definition. Two sequences of matrices $\mathbf{A}_n, \mathbf{B}_n$ are asymptotically equivalent, written $\mathbf{A}_n \simeq \mathbf{B}_n$, if for every sequence $\mathbf{\Theta}_n$ having uniformly bounded trace norm, almost surely

 $\lim_{n\to\infty} \operatorname{tr} \left[\boldsymbol{\Theta}_n (\mathbf{A}_n - \mathbf{B}_n) \right] = 0.$

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 $\lim_{n\to\infty} \operatorname{tr} \left[\mathbf{\Theta}_n (\mathbf{A}_n - \mathbf{B}_n) \right] = 0.$

• Admits a calculus:

- Addition $\mathbf{A}_n \simeq \mathbf{B}_n, \ \mathbf{C}_n \simeq \mathbf{D}_n \implies \mathbf{A}_n + \mathbf{C}_n \simeq \mathbf{B}_n + \mathbf{D}_n$
- Multiplication $\mathbf{A}_n \simeq \mathbf{B}_n \implies \mathbf{C}_n \mathbf{A}_n \mathbf{D}_n \simeq \mathbf{C}_n \mathbf{B}_n \mathbf{D}_n$
- Elements $\mathbf{A}_n \simeq \mathbf{B}_n \implies [\mathbf{A}_n]_{ij} [\mathbf{B}_n]_{ij} \xrightarrow{\text{a.s.}} 0$
- Differentiation [7] $f(\mathbf{A}_n; z) \simeq g(\mathbf{B}_n; z) \implies f'(\mathbf{A}_n; z) \simeq g'(\mathbf{B}_n; z)$

[6] E Dobriban, Y Sheng. "Distributed linear regression by averaging." Annals of Statistics, 2021.[7] E Dobriban, Y Sheng. "WONDER: Weighted one-shot distributed ridge regression in high dimensions." JMLR, 2020.

An Asymptotic Equivalence of Resolvents

Theorem (Rubio & Mestre 2011). Let $\mathbf{Z} \in \mathbb{C}^{n \times p}$ be a random matrix consisting of i.i.d. random variables that have mean 0, variance 1, and finite absolute moment of order $8+\delta$ for some $\delta > 0$. Let $\mathbf{\Sigma} \in \mathbb{C}^{p \times p}$ be a positive semidefinite matrix with operator norm uniformly bounded in p, and let $\mathbf{X} = \mathbf{Z}\mathbf{\Sigma}^{1/2}$. Then, for $z \in \mathbb{C}^+$, as $n, p \to \infty$ such that $0 < \liminf \frac{p}{n} \leq \lim \sup \frac{p}{n} < \infty$, we have

$$\left(\frac{1}{n}\mathbf{X}^{\mathsf{H}}\mathbf{X} - z\mathbf{I}_{p}\right)^{-1} \simeq \left(c(z)\boldsymbol{\Sigma} - z\mathbf{I}_{p}\right)^{-1},\tag{32}$$

where c(z) is the unique solution in \mathbb{C}^- to the fixed point equation

$$\frac{1}{c(z)} - 1 = \frac{1}{n} \operatorname{tr} \left[\mathbf{\Sigma} \left(c(z) \mathbf{\Sigma} - z \mathbf{I}_p \right)^{-1} \right].$$
(33)

Furthermore, $\frac{1}{p} \operatorname{tr} \left[\mathbf{\Sigma}(c(z)\mathbf{\Sigma} - z\mathbf{I}_p)^{-1} \right]$ is a Stieltjes transform of a certain positive measure on $\mathbb{R}_{\geq 0}$ with total mass $\frac{1}{p} \operatorname{tr}[\mathbf{\Sigma}]$.

[8] F Rubio, X Mestre. "Spectral convergence for a general class of random matrices." Statistics & Probability Letters, 2011.

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$$\left(\frac{1}{n}\mathbf{X}^{\mathsf{H}}\mathbf{X} - z\mathbf{I}_{p}\right)^{-1} \simeq (c(z)\boldsymbol{\Sigma} - z)$$
 What about real arguments?

where c(z) is the unique solution in \mathbb{C}^- to the fixed point equation

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Real-valued Asymptotic Equivalence

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$$\frac{1}{n}\mathbf{X}^{\mathsf{H}}\mathbf{X} - z\mathbf{I}_{p}\right)^{-1} \simeq (c(z)\boldsymbol{\Sigma} - z\mathbf{I}_{p})^{-1}, \qquad (32)$$

where c(z) is the unique solution in \mathbb{C}^- to the fixed point equation

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<u>Real-valued Asymptotic Equivalence</u>

Theorem 3. Let $\zeta_0, z_0 \in \mathbb{R}$ be the unique solutions, satisfying $\zeta_0 < \lambda_{\min}^+(\Sigma)$, to system of equations

$$1 = \frac{1}{n} \operatorname{tr} \left[\boldsymbol{\Sigma}^2 \left(\boldsymbol{\Sigma} - \zeta_0 \mathbf{I}_p \right)^{-2} \right], \quad z_0 = \zeta_0 \left(1 - \frac{1}{n} \operatorname{tr} \left[\boldsymbol{\Sigma} \left(\boldsymbol{\Sigma} - \zeta_0 \mathbf{I}_p \right)^{-1} \right] \right). \tag{34}$$

Then, for each $z \in \mathbb{R}$ satisfying $z < \liminf z_0$, as $n, p \to \infty$ such that $0 < \liminf \frac{p}{n} \leq 1$ $\limsup \frac{p}{n} < \infty$, we have

$$z\left(\frac{1}{n}\mathbf{X}^{\mathsf{H}}\mathbf{X} - z\mathbf{I}_{p}\right)^{-1} \simeq \zeta\left(\mathbf{\Sigma} - \zeta\mathbf{I}_{p}\right)^{-1},\tag{35}$$

where c(z) is the u

some $\delta > 0$. Let $\Sigma \in \mathbb{C}^{p \times p}$ be a positive semidefinite matrix with operator norm uniformly bounded in p, and let $\mathbf{X} = \mathbf{Z} \mathbf{\Sigma}^{1/2}$. Then, for $z \in \mathbb{C}^+$, as $n, p \to \infty$ such that $0 < \liminf \frac{p}{n} \leq \infty$ $\limsup \frac{p}{p} < \infty$, we have $\left(\frac{1}{n}\mathbf{X}^{\mathsf{H}}\mathbf{X} - z\mathbf{I}_{p}\right)$

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Analytic continuation

Furthermore, $\frac{1}{n} \operatorname{tr} \left[\mathbf{\Sigma} (c(z) \mathbf{\Sigma} - z \mathbf{I}_p)^{-1} \right]$ is a Stieltjes transform of a certain positive measure on $\mathbb{R}_{\geq 0}$ with total mass $\frac{1}{n} \operatorname{tr}[\Sigma]$.

Where $\zeta \in \mathbb{R}$ is the unique solution in $(-\infty, \zeta_0)$ to the fixed-point equation

$$z = \zeta \left(1 - \frac{1}{n} \operatorname{tr} \left[\mathbf{\Sigma} \left(\mathbf{\Sigma} - \zeta \mathbf{I}_p \right)^{-1} \right] \right).$$
(36)

Furthermore, as $n, p \to \infty$, $|\zeta + \frac{1}{v(z)}| \xrightarrow{\text{a.s.}} 0$, where v(z) is the companion Stieltjes transform of the spectrum of $\frac{1}{n}\mathbf{X}^{\mathsf{H}}\mathbf{X}$ given by

$$v(z) = \frac{1}{n} \operatorname{tr}\left[\left(\frac{1}{n} \mathbf{X} \mathbf{X}^{\mathsf{H}} - z \mathbf{I}_{n}\right)^{-1}\right],$$

and $|z_0 - \lambda_{\min}^+(\frac{1}{n}\mathbf{X}^{\mathsf{H}}\mathbf{X})| \xrightarrow{\text{a.s.}} 0.$

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 $1 = \frac{1}{n} \operatorname{tr} \left[\boldsymbol{\Sigma}^2 \left(\boldsymbol{\Sigma} - \zeta_0 \mathbf{I}_p \right)^{-2} \right], \quad z_0 = \zeta_0 \left(1 - \frac{1}{n} \operatorname{tr} \left[\boldsymbol{\Sigma} \left(\boldsymbol{\Sigma} - \zeta_0 \mathbf{I}_p \right)^{-1} \right] \right).$

Limits of negative regularization

Theorem (Rubio & Mestre 2011). Let $\mathbf{Z} \in \mathbb{C}^{n \times p}$ be a random matrix consisting of i.i.d. random variables that have mean 0, variance 1, and finite absolute moment of order $8 + \delta$ for some $\delta > 0$. Let $\mathbf{\Sigma} \in \mathbb{C}^{p \times p}$ be a positive semidefinite matrix with operator norm uniformly bounded in p, and let $\mathbf{X} = \mathbf{Z} \mathbf{\Sigma}^{1/2}$. Then, for $z \in \mathbb{C}^+$, as $n, p \to \infty$ such that $0 < \liminf_n \frac{p}{n} \leq 1$ lim sup $\frac{p}{n} < \infty$, we have

$$\left(\frac{1}{n}\mathbf{X}^{\mathsf{H}}\mathbf{X} - z\mathbf{I}_{p}\right)^{-1} \simeq \left(c(z)\boldsymbol{\Sigma} - z\mathbf{I}_{p}\right)^{-1},$$

where c(z) is the unique solution in \mathbb{C}^- to the fixed point equation

$$\frac{1}{c(z)} - 1 = \frac{1}{n} \operatorname{tr} \left[\boldsymbol{\Sigma} \left(c(z) \boldsymbol{\Sigma} - z \mathbf{I}_p \right)^{-1} \right].$$
(33)

(32)

Furthermore, $\frac{1}{p}$ tr $[\mathbf{\Sigma}(c(z)\mathbf{\Sigma} - z\mathbf{I}_p)^{-1}]$ is a Stieltjes transform of a certain positive measure on $\mathbb{R}_{\geq 0}$ with total mass $\frac{1}{p}$ tr $[\mathbf{\Sigma}]$.

Then, for each
$$z \in \mathbb{R}$$
 satisfying $z < \liminf z_0$, as $n, p \to \infty$ such that $0 < \liminf \frac{p}{n} \le n \sup \frac{p}{n} < \infty$, we have

$$z\left(\frac{1}{n}\mathbf{X}^{\mathsf{H}}\mathbf{X} - z\mathbf{I}_{p}\right)^{-1} \simeq \zeta\left(\mathbf{\Sigma} - \zeta\mathbf{I}_{p}\right)^{-1},\tag{35}$$

where $\zeta \in \mathbb{R}$ is the unique solution in $(-\infty, \zeta_0)$ to the fixed-point equation

$$z = \zeta \left(1 - \frac{1}{n} \operatorname{tr} \left[\mathbf{\Sigma} \left(\mathbf{\Sigma} - \zeta \mathbf{I}_p \right)^{-1} \right] \right).$$
(36)

Furthermore, as $n, p \to \infty$, $|\zeta + \frac{1}{v(z)}| \xrightarrow{\text{a.s.}} 0$, where v(z) is the companion Stieltjes transform of the spectrum of $\frac{1}{n} \mathbf{X}^{\mathsf{H}} \mathbf{X}$ given by

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and $|z_0 - \lambda_{\min}^+(\frac{1}{n}\mathbf{X}^\mathsf{H}\mathbf{X})| \xrightarrow{\text{a.s.}} 0.$

(34)

Real-valued Asymptotic Equivalence

Theorem 3. Let $\zeta_0, z_0 \in \mathbb{R}$ be the unique solutions, satisfying $\zeta_0 < \lambda_{\min}^+(\Sigma)$, to system of equations

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Reparameterization $f = \frac{1}{n} \mathbf{X}^{\mathsf{H}} \mathbf{X}$ given by $(\zeta + \frac{1}{v(z)}) \xrightarrow{\text{a.s.}} 0$, where v(z) is the companion Stieltjes transform

$$v(z) = rac{1}{n} \mathrm{tr} \Big[\left(rac{1}{n} \mathbf{X} \mathbf{X}^{\mathsf{H}} - z \mathbf{I}_n
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(33)

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Obtaining a Sketching Equivalence

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$$\mathbf{X} = \sqrt{q} \mathbf{S}^{\mathsf{H}} \mathbf{A}^{1/2}$$
$$\lambda = -z$$
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$$\mathbf{I}_{p} - \lambda \left(\mathbf{A}^{1/2} \mathbf{S} \mathbf{S}^{\mathsf{H}} \mathbf{A}^{1/2} + \lambda \mathbf{I}_{p} \right)^{-1} \simeq \mathbf{I}_{p} - \mu \left(\mathbf{A} + \mu \mathbf{I}_{p} \right)^{-1}$$
$$\implies \mathbf{A}^{1/2} \mathbf{S} \left(\mathbf{S}^{\mathsf{H}} \mathbf{A} \mathbf{S} + \lambda \mathbf{I}_{q} \right)^{-1} \mathbf{S}^{\mathsf{H}} \mathbf{A}^{1/2} \simeq \mathbf{A}^{1/2} \left(\mathbf{A} + \mu \mathbf{I}_{p} \right)^{-1} \mathbf{A}^{1/2}$$

<u>Obtaining a Sketching Equivalence</u>

<u>First-order Sketching Equivalence</u>

Theorem 4. For each $\lambda > \limsup \lambda_0$, as $q, p \to \infty$,

$$\mathbf{S} (\mathbf{S}^{\mathsf{H}} \mathbf{A} \mathbf{S} + \lambda \mathbf{I}_q)^{-1} \mathbf{S}^{\mathsf{H}} \simeq (\mathbf{A} + \mu \mathbf{I}_p)^{-1},$$

where μ is the unique solution in (μ_0, ∞) to the fixed point equation

$$\lambda = \mu \left(1 - \frac{1}{q} \operatorname{tr} \left[\mathbf{A} \left(\mathbf{A} + \mu \mathbf{I}_p \right)^{-1} \right] \right).$$

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• That is, sketching + ridge = another ridge without sketching.



<u>First-order Sketching Equivalence</u>

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is the unique solution in (μ_{0}, ∞) to the fixed point $\mathbb{E} \left[\frac{q_{\mathbf{S}}}{p} \right] = \operatorname{tr} \left(\mathbf{M} \left(\mathbf{M} + \mu \mathbf{I} \right)^{-1} \right)$
 $\lambda = \mu \left(1 - \frac{1}{q} \operatorname{tr} \left[\mathbf{A} \left(\mathbf{A} + \mu \mathbf{I}_{p} \right)^{-1} \right] \right).$
Same as DPP when $\lambda = 0$

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• That is, sketching +

Spoiler:

ketching.

Randomized least squares = ridge + noise

Randomized ensembles = ridge



<u>Element-wise Convergence</u>



Element-wise Convergence








16x16 🛛



16x16 🔤 🔤







• Example: isotropic spectrum with $r(\mathbf{A}) = \frac{1}{2}$



- Example: isotropic spectrum with $r(\mathbf{A}) = \frac{1}{2}$
- ${\:}^{\bullet}\lambda\mapsto\mu$ is increasing and concave
- $\bullet \alpha \mapsto \mu \text{ is decreasing unless } \alpha > r(\mathbf{A}) \text{ and } \lambda < 0$



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•
$$\mu \leq \lambda + \frac{1}{q} \operatorname{tr}[\mathbf{A}]$$



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- $\mu \leq \lambda + \frac{1}{q} \operatorname{tr}[\mathbf{A}]$ • $\operatorname{sign}(\mu) = \operatorname{sign}(\lambda)$ if $\alpha > r(\mathbf{A})$ • else $\mu \geq 0$



First-Order is Not Enough

• First-order equivalence is similar to expectation equivalence

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- $\mathbb{E}[X] = \mathbb{E}[Y]$ does not imply that $\mathbb{E}[X^2] = \mathbb{E}[Y^2]$
- Similarly, products of equivalences do not compose if not independent

- First-order equivalence is similar to expectation equivalence
- $\mathbb{E}[X] = \mathbb{E}[Y]$ does not imply that $\mathbb{E}[X^2] = \mathbb{E}[Y^2]$
- Similarly, products of equivalences do not compose if not independent
- Solution: derivative rule of asymptotic equivalence

$$\frac{d}{dz} \left(\mathbf{A} - z\mathbf{I} \right)^{-1} = -\left(\mathbf{A} - z\mathbf{I} \right)^{-2}$$

Second-order Sketching Equivalence

Theorem 5. If $\Psi \in \mathbb{C}^{p \times p}$ is a deterministic or random positive semidefinite matrix independent of **S** with $\|\Psi\|_{op}$ uniformly bounded in p, then

$$\mathbf{S} \left(\mathbf{S}^{\mathsf{H}} \mathbf{A} \mathbf{S} + \lambda \mathbf{I}_{q} \right)^{-1} \mathbf{S}^{\mathsf{H}} \Psi \mathbf{S} \left(\mathbf{S}^{\mathsf{H}} \mathbf{A} \mathbf{S} + \lambda \mathbf{I}_{q} \right)^{-1} \mathbf{S}^{\mathsf{H}} \simeq \left(\mathbf{A} + \mu \mathbf{I}_{p} \right)^{-1} \left(\Psi + \mu' \mathbf{I}_{p} \right) \left(\mathbf{A} + \mu \mathbf{I}_{p} \right)^{-1},$$

$$\mu' = \frac{\frac{1}{q} \operatorname{tr} \left[\mu^3 \left(\mathbf{A} + \mu \mathbf{I}_p \right)^{-1} \mathbf{\Psi} \left(\mathbf{A} + \mu \mathbf{I}_p \right)^{-1} \right]}{\lambda + \frac{1}{q} \operatorname{tr} \left[\mu^2 \mathbf{A} \left(\mathbf{A} + \mu \mathbf{I}_p \right)^{-2} \right]} \ge 0.$$

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where
$$\frac{1}{q} \operatorname{tr} \left[\mu^{3} \left(\mathbf{A} + \mu \mathbf{I}_{p} \right)^{-1} \Psi \left(\mathbf{A} + \mu \mathbf{I}_{p} \right)^{-1} \right] > 0$$

wh

$$\mu' = \frac{\frac{1}{q} \operatorname{tr} \left[\mu^3 \left(\mathbf{A} + \mu \mathbf{I}_p \right)^{-1} \mathbf{\Psi} \left(\mathbf{A} + \mu \mathbf{I}_p \right)^{-1} \right]}{\lambda + \frac{1}{q} \operatorname{tr} \left[\mu^2 \mathbf{A} \left(\mathbf{A} + \mu \mathbf{I}_p \right)^{-2} \right]} \ge 0.$$

- That is, second-order adds an isotropic inflation factor
- Depends significantly on the choice of Ψ

Second-order Sketching Equivalence

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One Incredible Regime

• Example: isotropic spectrum with $r(\mathbf{A}) = \frac{1}{2}$



One Incredible Regime





- If $\Psi \in \operatorname{Range}(\mathbf{A}), \, \alpha > r(\mathbf{A}),$ and $\mu = \lambda = 0$, there is no inflation

- Agrees with classical sketching results: sketch larger than rank
- Sketching is ideal for benign overfitting

Application: Ridge Regression



<u>Sketched Ridge Regression</u>

 $\widehat{\boldsymbol{\beta}}_{\boldsymbol{P}} \triangleq \operatorname*{arg\,min}_{\mathbf{b}} \frac{1}{n} \left\| \mathbf{T}^{\top} (\mathbf{y} - \mathbf{X}\mathbf{b}) \right\|_{2}^{2} + \lambda \|\mathbf{b}\|_{2}^{2}$ $\widehat{\boldsymbol{\beta}}_{D} \triangleq \mathbf{S} \cdot \operatorname*{arg\,min}_{\mathbf{b}} \frac{1}{n} \|\mathbf{y} - \mathbf{X}\mathbf{S}\mathbf{b}\|_{2}^{2} + \lambda \|\mathbf{b}\|_{2}^{2}$ $\widehat{\boldsymbol{\beta}}_{PD} \triangleq \mathbf{S} \cdot \operatorname*{arg\,min}_{\mathbf{b}} \frac{1}{n} \|\mathbf{T}^{\top}(\mathbf{y} - \mathbf{X}\mathbf{S}\mathbf{b})\|_{2}^{2} + \lambda \|\mathbf{b}\|_{2}^{2}$

Primal (observations) and dual (features) sketching:

Preparing for Equivalences

• Express in terms of the sketched pseudoinverse:

$$\widehat{\boldsymbol{\beta}}_{\psi} = \frac{1}{\sqrt{n}} \mathbf{X}_{\psi}^{\ddagger} \mathbf{y}, \ \psi \in \{\boldsymbol{P}, \boldsymbol{D}, \boldsymbol{P}\boldsymbol{D}\}$$
$$\mathbf{X}_{\boldsymbol{P}}^{\ddagger} \triangleq \frac{1}{\sqrt{n}} \left(\frac{1}{n} \mathbf{X}^{\top} \mathbf{T} \mathbf{T}^{\top} \mathbf{X} + \lambda \mathbf{I}\right)^{-1} \mathbf{X}^{\top} \mathbf{T} \mathbf{T}^{\top}$$
$$\mathbf{X}_{\boldsymbol{D}}^{\ddagger} \triangleq \frac{1}{\sqrt{n}} \mathbf{S} \left(\frac{1}{n} \mathbf{S}^{\top} \mathbf{X}^{\top} \mathbf{X} \mathbf{S} + \lambda \mathbf{I}\right)^{-1} \mathbf{S}^{\top} \mathbf{X}^{\top}$$
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<u>First-order Data Pseudoinverse Equivalence</u>

Theorem 6. If $\left\|\frac{1}{\sqrt{n}}\mathbf{X}\right\|_{\text{op}}$ is uniformly bounded in p, then as $m, n, q, p \to \infty$,

$$\mathbf{X}_{\psi}^{\ddagger} \simeq \frac{1}{\sqrt{n}} \left(\frac{1}{n} \mathbf{X}^{\mathsf{H}} \mathbf{X} + \mu_{\psi} \mathbf{I} \right)^{-1} \mathbf{X}^{\mathsf{H}},$$

where μ_{ψ} are is the most positive solutions to the equations

$$\lambda = \mu_{P} \left(1 - \frac{1}{m} \operatorname{tr} \left[\frac{1}{n} \mathbf{X}^{\mathsf{H}} \mathbf{X} \left(\frac{1}{n} \mathbf{X}^{\mathsf{H}} \mathbf{X} + \mu_{P} \mathbf{I} \right)^{-1} \right] \right),$$
$$\lambda = \mu_{D} \left(1 - \frac{1}{q} \operatorname{tr} \left[\frac{1}{n} \mathbf{X}^{\mathsf{H}} \mathbf{X} \left(\frac{1}{n} \mathbf{X}^{\mathsf{H}} \mathbf{X} + \mu_{D} \mathbf{I} \right)^{-1} \right] \right),$$
$$\frac{\lambda}{\theta} - 1 = \frac{m}{q} \left(\frac{\theta}{\mu_{PD}} - 1 \right), \quad \theta = \mu_{PD} \left(1 - \frac{1}{m} \operatorname{tr} \left[\frac{1}{n} \mathbf{X}^{\mathsf{H}} \mathbf{X} \left(\frac{1}{n} \mathbf{X}^{\mathsf{H}} \mathbf{X} + \mu_{PD} \mathbf{I} \right)^{-1} \right] \right).$$





<u>Sketching Makes Same Predictions as Ridge</u>

Corollary 7. If $\|\frac{1}{\sqrt{n}}\mathbf{y}\|_2$ is uniformly bounded in p and $\mathbf{w} \in \mathbb{C}^p$ is independent of \mathbf{S} and \mathbf{T} such that $\|\mathbf{w}\|_2$ is uniformly bounded in p, then for any continuous function $f: \mathbb{C} \to \mathbb{C}$, as $p \to \infty$,

$$f(\mathbf{w}^{\mathsf{H}}\widehat{\boldsymbol{\beta}}_{\psi}) - f(\mathbf{w}^{\mathsf{H}}\widehat{\boldsymbol{\beta}}_{\mu_{\psi}}^{\mathrm{ridge}}) \xrightarrow{\mathrm{a.s.}} 0.$$

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Sketching Makes Same Predictions as Ridge

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• Sketching asymptotically makes the same prediction as the equivalent ridge on any single test point Why this qualifier?

Sketching Makes Same Predictions as Ridge

Corollary 7. If $\|\frac{1}{\sqrt{n}}\mathbf{y}\|_2$ is uniformly bounded in p and $\mathbf{w} \in \mathbb{C}^p$ is independent of \mathbf{v} and ${f T}$ such that $\|\mathbf{w}\|_2$ is uniformly bounded in p, then for any continuous funct \mathbb{C}, as $p \to \infty$,

Pointwise convergence does not imply uniform convergence prediction as the equivalent Sketching asymptotic ridge on any sin

- Ensemble of independent sketches: $\hat{\beta}_{\psi}^{\text{ens}} = \frac{1}{K} \sum_{k=1}^{K} \hat{\beta}_{\psi}^{(k)}$
- Quadratic error metrics: $\mathcal{E}_{\Psi}(\boldsymbol{\beta}, \boldsymbol{\beta}') = (\boldsymbol{\beta} \boldsymbol{\beta}')^{\mathsf{H}} \Psi (\boldsymbol{\beta} \boldsymbol{\beta}')$
 - Includes test risk

Theorem 8. If $\Psi \in \mathbb{C}^{p \times p}$ is a positive semidefinite matrix and $\beta' \in \mathbb{C}^p$ a vector such that $\|\Psi\|_{\text{op}}$ and $\|\beta'\|_2$ are uniformly bounded in p and (Ψ, β) is independent of $(\mathbf{S}_k, \mathbf{T}_k)_{k=1}^K$, then for $\psi \in \{P, D\}$,

$$\mathcal{E}_{\Psi}(\widehat{\boldsymbol{\beta}}_{P}^{\text{ens}},\boldsymbol{\beta}') - \left(\mathcal{E}_{\Psi}(\widehat{\boldsymbol{\beta}}_{\mu_{P}}^{\text{ridge}},\boldsymbol{\beta}') + \frac{\mu_{P}'}{Kn}\mathbf{y}^{\mathsf{H}}\left(\frac{1}{n}\mathbf{X}\mathbf{X}^{\mathsf{H}} + \mu_{P}\mathbf{I}\right)^{-2}\mathbf{y}\right) \xrightarrow{\text{a.s.}} 0, \\ \mathcal{E}_{\Psi}(\widehat{\boldsymbol{\beta}}_{D}^{\text{ens}},\boldsymbol{\beta}') - \left(\mathcal{E}_{\Psi}(\widehat{\boldsymbol{\beta}}_{\mu_{D}}^{\text{ridge}},\boldsymbol{\beta}') + \frac{\mu_{D}'}{Kn}\mathbf{y}^{\mathsf{H}}\mathbf{X}\left(\frac{1}{n}\mathbf{X}^{\mathsf{H}}\mathbf{X} + \mu_{D}\mathbf{I}\right)^{-2}\mathbf{X}^{\mathsf{H}}\mathbf{y}\right) \xrightarrow{\text{a.s.}} 0,$$

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$$\text{Well, ain't this a geometrical oddity.}$$

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$$\mathcal{E}_{\Psi}(\widehat{\boldsymbol{\beta}}_{\psi}^{\text{ens}}, \boldsymbol{\beta}') \xrightarrow{\text{a.s.}} \mathcal{E}_{\Psi}(\widehat{\boldsymbol{\beta}}_{\mu_{\psi}}^{\text{ridge}}, \boldsymbol{\beta}') + O\left(\frac{\mu_{\psi}'}{K}\right)$$

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Spoiler:

Randomized least squares = ridge + noise

Randomized ensembles = ridge

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- For infinite *K*, sketched ensemble = ridge
- For finite K, sketched ensemble is worse than ridge
 - Unless $\Psi \in \operatorname{Range}(\mathbf{A})$, $\alpha > r(\mathbf{A})$, and $\mu = \lambda = 0$!


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A Sketched Ensemble Efficiency Experiment • Setup: fixed $O(p^2n)$ budget ensembles with $K = \lfloor \frac{1}{\alpha^2} \rfloor$, $r(\Sigma) = \frac{1}{2}$ • Fixed target μ_D , varying lpha, with λ uniquely determined • Error: relative error $\underline{\mathcal{E}_{\Psi}(\widehat{\boldsymbol{\beta}}_{D}^{\text{ens}},\widehat{\boldsymbol{\beta}}_{\mu_{D}}^{\text{ridge}})}$ $\mathcal{E}_{oldsymbol{\Psi}}(oldsymbol{0},\widehat{oldsymbol{eta}}_{\mu_{\mathcal{D}}}^{\mathrm{ridge}})$ Non-vanishing estimation error Estimation $(\Psi = \mathbf{I}_p)$ Prediction ($\Psi = \Sigma$) μ_D 10^{4} 10^{4} - 10 Relative Error 10^2 10^2 0.1 10^0 10^{0} 0.01 10^{-2} 10^{-2} -0.01 10^{-4} 10^{-4} 0.30.50.60.20.3 0.40.50.20.40.70.60.70.10.8 0.10.8 α α

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Better Sketches?



$$\mathbf{S} \left(\mathbf{S}^{\mathsf{H}} \mathbf{A} \mathbf{S} + \lambda \mathbf{I}_{q} \right)^{-1} \mathbf{S}^{\mathsf{H}} \simeq \left(\mathbf{A} + \gamma \mathbf{I}_{p} \right)^{-1}.$$

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 - Orthogonal sketching
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Orthogonal Sketching

Conjecture 10. For $q \leq p$ let $\sqrt{\frac{q}{p}} \mathbf{Q} \in \mathbb{C}^{p \times q}$ be a Haar-distributed matrix with orthonormal columns. Then

$$\mathbf{Q} (\mathbf{Q}^{\mathsf{H}} \mathbf{A} \mathbf{Q} + \lambda \mathbf{I}_{q})^{-1} \mathbf{Q}^{\mathsf{H}} \simeq (\mathbf{A} + \gamma \mathbf{I}_{p})^{-1},$$

where γ is the most positive solution to

$$\frac{1}{p} \operatorname{tr} \left[\left(\mathbf{A} + \gamma \mathbf{I}_p \right)^{-1} \right] \left(\gamma - \alpha \lambda \right) = 1 - \alpha.$$

Furthermore, for $\mu > 0$ applied to the same $(\mathbf{A}, \alpha, \lambda)$, we have $\gamma < \mu$.

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Furthermore, for $\mu > 0$ *applied to the same* (**A**, α , λ)*, we have* $\gamma < \mu$. • Same form as i.i.d. sketching, but with less regularization

Equivalence for Sketches Used in Practice?



• Early work:

• Hints of deep connection between ensembles and ridge

• Tuned ensembles with subsampling achieve same risk as optimal ridge

Sumi

marv

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• Future work:

- More asymptotic equivalences
 - Generalized cross-validation with sketching
 - General linear models via leave-one-dimension-out
 - Asymptotics of PCA



Questions?



Corollary 11. Let \mathbf{W} be an invertible $p \times p$ positive semidefinite matrix, either deterministic or random but independent of \mathbf{S} with $\limsup \|\mathbf{W}\|_{op} < \infty$. Let $\widetilde{\mathbf{S}} = \mathbf{W}^{1/2}\mathbf{S}$. Then for each $\lambda > -\liminf \lambda_{\min}^{+}(\widetilde{\mathbf{S}}^{\top}\mathbf{A}\widetilde{\mathbf{S}})$ as $p, q \to \infty$ such that $0 < \liminf \inf \frac{q}{p} \le \limsup \frac{q}{p} < \infty$,

$$\widetilde{\mathbf{S}} \big(\widetilde{\mathbf{S}}^{\top} \mathbf{A} \widetilde{\mathbf{S}} + \lambda \mathbf{I}_q \big)^{-1} \widetilde{\mathbf{S}}^{\top} \simeq \big(\mathbf{A} + \mu \mathbf{W}^{-1} \big)^{-1},$$

where μ most positive solution to

$$\lambda = \mu \left(1 - \frac{1}{q} \operatorname{tr} \left[\mathbf{A} \left(\mathbf{A} + \mu \mathbf{W}^{-1} \right)^{-1} \right] \right).$$

Theorem 12. Let $\mathbf{X} = \mathbf{Z} \mathbf{\Sigma}^{1/2}$ and $\mathbf{\Sigma} \in \mathbb{C}^{p \times p}$ have eigenvalue decomposition $\mathbf{U} \mathbf{D} \mathbf{U}^{\mathsf{H}}$, and let $\mathbf{\Pi}_{\mathcal{A}}$ be the projection operator of the principal eigenspace corresponding to a set of eigenvalues \mathcal{A} of the matrix $\frac{1}{n} \mathbf{X}^{\mathsf{H}} \mathbf{X}$. Then there exists a family of measures μ_{σ^2} for all $\sigma^2 \geq 0$ such that for any $\mathcal{A} \subseteq \mathbb{R}_{\geq 0}$, in the limit as $p \to \infty$,

 $\Pi_{\mathcal{A}}\simeq \mathbf{U} \boldsymbol{\Lambda} \mathbf{U}^{\mathsf{H}},$

where Λ is a diagonal matrix defined for by

 $[\mathbf{\Lambda}]_{ii} = \mu_{[\mathbf{\Sigma}]_{ii}}(\mathcal{A}).$