

# 1 More sub-Gaussian characterizations

Throughout,  $C, c > 0$  denote absolute constants (that may change from line to line), and we write

$$A \lesssim B \quad \text{to mean} \quad A \leq CB.$$

Let  $X$  be a real-valued random variable, and define the  $\psi_2$  (Orlicz) norm

$$\|X\|_{\psi_2} := \inf \left\{ K > 0 : \mathbb{E} \exp\left(\frac{X^2}{K^2}\right) \leq 2 \right\}.$$

Recall that  $X$  is sub-Gaussian if and only if  $\|X\|_{\psi_2} < \infty$ .

- (a) **Sub-Gaussianity via (almost) Gaussian tail domination.** Let  $g \sim \mathcal{N}(0, 1)$ . For two random variables  $U, V$ , define the relation

$$U \leq V \quad \iff \quad \mathbb{P}\{|U| \geq t\} \leq 2\mathbb{P}\{|V| \geq t\} \quad \text{for all } t \geq 0.$$

Show that  $X$  is sub-Gaussian if and only if there exists  $K > 0$  such that

$$X \leq Kg. \tag{1}$$

More precisely:

- (i) If  $\|X\|_{\psi_2} \leq L$ , prove that (1) holds with  $K \leq CL$ .

**Solution:**

Using Exercise 5 in HW2,  $\|X\|_{\psi_2} \leq L$  implies:

$$\mathbb{P}\{|X| \geq t\} \leq 2 \exp\left(-c_1 \frac{t^2}{L^2}\right), \quad \text{for all } t \geq 0. \tag{2}$$

Using Mill's ratio for all  $u > 0$ ,

$$\mathbb{P}\{g \geq u\} \geq \frac{\varphi(u)}{u + u^{-1}}, \quad \text{where} \quad \varphi(u) = \frac{1}{\sqrt{2\pi}} e^{-u^2/2}.$$

Because  $g$  is symmetric around 0,  $\mathbb{P}\{|g| \geq u\} = 2\mathbb{P}\{g \geq u\} \geq \frac{2u}{u^2 + 1} e^{-u^2/2}$ . So,

$$\mathbb{P}\{|g| \geq u\} \geq \frac{c_2}{u + 1} e^{-u^2/2} \quad \text{for all } u > \frac{1}{2}. \tag{3}$$

Using the crude inequality  $\log(u + 1) \leq u^2/4 + 1$  in (3) gives:

$$\mathbb{P}\{|g| \geq u\} \geq c_2 e^{-1} e^{-3u^2/4} \geq c_3 e^{-3u^2/4}, \quad (u > 1/2),$$

for some absolute constant  $c_3 > 0$ . Let  $u = t/K$ , which gives

$$2\mathbb{P}\left\{|g| \geq \frac{t}{K}\right\} \geq 2c_3 \exp\left(-\frac{3t^2}{4K^2}\right), \quad (t > K/2). \tag{4}$$

We need to show  $\mathbb{P}\{|X| \geq t\} \leq 2\mathbb{P}\{|g| \geq t/K\}$ , so using (2) and (4) it suffices to show:

$$\exp\left(-c_1 \frac{t^2}{L^2}\right) \leq c_3 \exp\left(-\frac{3t^2}{4K^2}\right) \iff \exp\left(-t^2 \left(\frac{c_1}{L^2} - \frac{3}{4K^2}\right)\right) \leq c_3. \tag{5}$$

First choose  $K = CL$  such that the coefficient in (5) is positive, i.e.,

$$\frac{c_1}{L^2} - \frac{3}{4K^2} = \frac{1}{L^2} \left( c_1 - \frac{3}{4C^2} \right) \geq 0 \iff C \geq \sqrt{\frac{3}{4c_1}}, \quad (6)$$

which is an absolute constant. Now, the exponential in (5) is a decreasing function of  $t$  when the coefficient is positive, so it's enough to satisfy the bound on the exponential on the left endpoint, which is for  $t = K/2$ . Putting  $t = K/2$  in (5) gives us the condition:

$$\exp\left(\frac{3}{16} - \frac{c_1 C^2}{4}\right) \leq c_3 \iff C^2 \geq \frac{3}{4c_1} - \frac{4 \log c_3}{c_1}. \quad (7)$$

Now, if  $u \in [0, 1/2]$  then  $t/K \leq 1/2$ . This gives  $\mathbb{P}\{|g| \geq t/K\} \geq \mathbb{P}\{|g| \geq 1/2\}$ , and  $\mathbb{P}\{|g| \geq 1/2\} > 1/2$  because for a standard normal,  $z_{0.25} = 0.674$ . So,

$$\mathbb{P}\{|g| \geq t/K\} > 1/2 \iff 2\mathbb{P}\{|g| \geq t/K\} > 1.$$

However, for any  $t$ ,  $\mathbb{P}\{|X| \geq t\} \leq 1$ , so the bound  $\mathbb{P}\{|X| \geq t\} \leq 2\mathbb{P}\{|g| \geq t/K\}$  holds trivially true. Hence, combining (6) and (7),  $X \leq Kg$  for all  $K \geq CL$  for absolute constant  $C$  given by:

$$C^2 = \max\left\{\frac{3}{4c_1}, \frac{3}{4c_1} - \frac{4 \log c_3}{c_1}\right\}. \quad \square$$

(ii) Conversely, if (1) holds for some  $K$ , prove that  $\|X\|_{\psi_2} \leq CK$ .

**Solution:**

Using the assumption that  $X \leq Kg$  for some  $K > 0$ ,

$$\begin{aligned} \mathbb{P}\{|X| \geq t\} &\leq 2\mathbb{P}\{|g| \geq t/K\} \\ &\leq 2 \cdot 2 \exp\left(-\frac{t^2}{2K^2}\right) \quad [g \text{ is gaussian, so has tail bound.}] \\ &= 2 \exp\left(\log 2 - \frac{t^2}{2K^2}\right) \\ &\leq 2 \exp\left(-\frac{t^2}{4K^2}\right), \end{aligned} \quad (8)$$

where the last inequality holds because  $\log 2 - \frac{t^2}{2K^2} \leq -\frac{t^2}{4K^2}$  whenever  $\frac{t^2}{4K^2} \geq \log 2$  and for smaller  $t$ ,  $2 \exp\left(-\frac{t^2}{4K^2}\right) \geq 2e^{-\log 2} = 1$  while probability is always  $\leq 1$ . Hence, (8) holds for all  $t \geq 0$ , and so  $X$  is  $(2K^2)$ -sub-Gaussian. Now, using the result from Exercise 5 in HW2, there exists an absolute constant  $c_1 > 0$  such that

$$\mathbb{E} \exp\left(\frac{X^2}{c_1 \cdot 2K^2}\right) \leq 2.$$

Hence, by the definition of  $\psi_2$ -norm,  $\|X\|_{\psi_2} \leq CK$  for some absolute constant  $C > 0$ .

- (b) [Bonus] Show by example that part (a) can fail if the definition of  $\leq$  is strengthened by removing the factor 2, i.e., if one demands  $\mathbb{P}\{|U| \geq t\} \leq \mathbb{P}\{|V| \geq t\}$  for all  $t \geq 0$ .

**[Bonus] Solution:**

Suppose we strengthen the relation by defining

$$X \text{ is sub-Gaussian} \iff \mathbb{P}\{|X| \geq t\} \leq \mathbb{P}\{|g| \geq t/K\}, \quad \text{for all } t \geq 0, \quad (9)$$

for some  $K > 0$ . Let  $X = \varepsilon$ , where  $\varepsilon \in \{-1, +1\}$  be a Rademacher random variable. We know,  $X$  is sub-Gaussian. Now, for any  $t \in (0, 1]$ ,  $\mathbb{P}\{|X| \geq t\} = 1$  because  $|X| = 1$  a.s. Pick  $t = 1/2$ , then for any  $K > 0$ ,  $\mathbb{P}\left\{|g| \geq \frac{1}{2K}\right\} < 1$  because a 1D-Gaussian has full support on  $\mathbb{R}$  and so  $\mathbb{P}\{|g| \geq u\} < 1$  for every  $u > 0$ . Hence,

$$\mathbb{P}\left\{|X| \geq \frac{1}{2}\right\} = 1 \not\leq \mathbb{P}\left\{|g| \geq \frac{1}{2K}\right\},$$

for any finite  $K > 0$ , and definition (9) fails. Now, if we let the original relation:

$$X \text{ is sub-Gaussian} \iff \mathbb{P}\{|X| \geq t\} \leq 2\mathbb{P}\{|g| \geq t/K\}, \quad \text{for all } t \geq 0, \quad (10)$$

for some  $K > 0$ , then we only need to check it at  $t = 1$ . We need, for  $t = 1$ ,

$$\mathbb{P}\{|X| \geq 1\} = 1 \leq 2\mathbb{P}\left\{|g| \geq \frac{1}{K}\right\},$$

which can be achieved by choosing  $K$  such that  $\mathbb{P}\left\{|g| \geq \frac{1}{K}\right\} \geq \frac{1}{2}$ , and (10) holds.  $\square$

- (c) [Bonus] **A local MGF of  $X^2$  characterization.** Show that  $X$  is sub-Gaussian if and only if there exists  $K > 0$  such that

$$\mathbb{E} \exp(\lambda^2 X^2) \leq \exp(\lambda^2 K^2) \quad \text{for all } |\lambda| \leq \frac{1}{K}. \quad (11)$$

More precisely:

- (i) Assume  $\|X\|_{\psi_2} \leq L$ . Prove that (11) holds with  $K \leq CL$ .

**[Bonus] Solution:**

Define  $U := \frac{X^2}{L^2}$ , then using  $\|X\|_{\psi_2} \leq L$ ,  $\mathbb{E}e^U = \mathbb{E} \exp\left(\frac{X^2}{L^2}\right) \leq 2$ . Let  $\theta = \lambda^2 L^2$  so that  $\theta \leq 1$  when  $|\lambda| \leq 1/L$ . For  $\theta \in [0, 1]$ ,  $x \mapsto x^\theta$  is concave on  $\mathbb{R}^+$ , so using Jensen's inequality for concave functions,

$$\mathbb{E} \exp(\lambda^2 X^2) = \mathbb{E}e^{U\theta} = \mathbb{E}(e^U)^\theta \leq (\mathbb{E}e^U)^\theta \leq 2^\theta = \exp((\log 2)\lambda^2 L^2) \leq \exp(\lambda^2 K^2). \quad (12)$$

Now, to prove (11), it suffices to choose  $K$  such that

$$(\log 2)\lambda^2 L^2 \leq \lambda^2 K^2 \iff K \geq \sqrt{\log 2} L.$$

We also need  $\{\lambda : |\lambda| \leq 1/K\} \subseteq \{\lambda : |\lambda| \leq 1/L\}$  so we can use (12), which holds if  $K \geq L$ .

Thus, choose  $K = L$  as  $\log 2 < 1$ . Then,

$$\mathbb{E} \exp(\lambda^2 X^2) \leq \exp((\log 2)\lambda^2 L^2) \leq \exp(\lambda^2 L^2) = \exp(\lambda^2 K^2), \quad \text{for all } |\lambda| \leq \frac{1}{K}.$$

Hence, (11) holds with  $K = L \leq CL$  for  $C = 1$ , and also holds for all  $K' \geq K$ .  $\square$

(ii) Conversely, assume (11) holds for some  $K > 0$ . Prove that  $X$  is sub-Gaussian with  $\|X\|_{\psi_2} \leq 2K$ .

**[Bonus] Solution:**

Here, (11) holds for some  $K > 0$ , i.e., for some  $K > 0$ ,

$$\mathbb{E} \exp(\lambda^2 X^2) \leq \exp(\lambda^2 K^2) \quad \text{for all } |\lambda| \leq \frac{1}{K}.$$

Substitute  $\lambda = \frac{1}{2K} \leq \frac{1}{K}$ , which gives

$$\mathbb{E} \exp\left(\frac{X^2}{(2K)^2}\right) \leq \exp\left(\frac{1}{4K^2} \cdot K^2\right) = \exp\left(\frac{1}{4}\right) < 2.$$

Hence, by the definition of  $\psi_2$ -norm,  $\|X\|_{\psi_2} \leq 2K$  and  $X$  is sub-Gaussian.  $\square$

## 2 Sub-exponential characterizations

Let  $Z$  be a real-valued random variable with  $\mathbb{E}Z = 0$ . Define the centered log-MGF

$$\psi_Z(\lambda) := \log \mathbb{E}e^{\lambda Z} \in (-\infty, \infty] \quad (\lambda \in \mathbb{R}),$$

and define the  $\psi_1$  (Orlicz) norm

$$\|X\|_{\psi_1} := \inf \left\{ K > 0 : \mathbb{E} \exp\left(\frac{|Z|}{K}\right) \leq 2 \right\}.$$

(a) **Four sub-exponential properties.** Fix  $K > 0$ . Consider the statements:

(i) Tail bound: For all  $t \geq 0$ ,

$$\mathbb{P}\{|Z| \geq t\} \leq 2 \exp\left(-c \frac{t}{K}\right).$$

(ii) Moment growth: There exists an absolute constant  $C > 0$  such that for all  $p \geq 1$ ,

$$(\mathbb{E}|Z|^p)^{1/p} \leq CKp.$$

(iii)  $\psi_1$ -integrability:

$$\mathbb{E} \exp\left(\frac{|Z|}{K}\right) \leq 2.$$

(iv) Local quadratic log-MGF: There exist absolute constants  $c, C > 0$  such that for all  $|\lambda| \leq \frac{c}{K}$ ,

$$\psi_Z(\lambda) \leq C\lambda^2 K^2.$$

Our goal is to show that (i)–(iv) are equivalent up to universal constants (i.e., the parameter  $K$  may change by at most an absolute constant factor from one statement to another).

(a1) (i)  $\implies$  (ii).

### Solution:

From Exercise 4 (a2), in HW1, for a non-negative random variable  $X$ , if there exist constants  $c, C > 0$  and  $\alpha > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P}\{X \geq t\} \leq C \exp(-ct^\alpha),$$

then there exists a constant  $C' = C'(C, \alpha)$  such that

$$\|X\|_{L_p} = (\mathbb{E}X^p)^{1/p} \leq C'(p/c)^{1/\alpha}, \quad \text{for all } p \geq 1. \quad (13)$$

For  $X = |Z|$ ,  $\alpha = 1$ , and  $c = 1/K$ , using (13), the tail bound,

$$\mathbb{P}\{|Z| \geq t\} \leq 2 \exp\left(-c \frac{t}{K}\right),$$

for all  $t \geq 0$  implies that there exists an absolute constant  $C > 0$  such that

$$(\mathbb{E}|Z|^p)^{1/p} \leq CKp. \quad \text{for all } p \geq 1. \quad \square \quad (14)$$

(a2) (ii)  $\implies$  (iii).**Solution:**

Using the power series expansion of exponential,  $e^x = \sum_{n=0}^{\infty} \frac{x^n}{n!}$ , gives,

$$\begin{aligned} \mathbb{E} \exp\left(\frac{|Z|}{K}\right) &= 1 + \sum_{n=1}^{\infty} \frac{\mathbb{E}|Z|^n}{K^n n!} \leq 1 + \sum_{n=1}^{\infty} \frac{\mathbb{E}|Z|^n}{K^n (n/e)^n} && [\because n! \geq (n/e)^n] \\ &\leq 1 + \sum_{n=1}^{\infty} \frac{C^n K_1^n n^n}{K^n (n/e)^n} && [\because \text{moment bound (14)}] \\ &= 1 + \sum_{n=1}^{\infty} \left(\frac{eCK_1}{K}\right)^n. \end{aligned}$$

Now, if we choose  $K \geq 2eCK_1$ , then

$$\mathbb{E} \exp\left(\frac{|Z|}{K}\right) \leq 1 + \sum_{n=1}^{\infty} 2^{-n} = 1 + 1 = 2.$$

Hence, moment growth in (ii) implies  $\psi_1$ -integrability, i.e.,

$$\mathbb{E} \exp\left(\frac{|Z|}{K}\right) \leq 2. \quad \square \tag{15}$$

(a3) (iii)  $\implies$  (iv).**Solution:**

Using the Taylor expansion,  $e^x = 1 + x + \sum_{m=2}^{\infty} \frac{x^m}{m!}$ , for  $x = \lambda Z$  gives  $e^{\lambda Z} = 1 + \lambda Z + \sum_{m=2}^{\infty} \frac{\lambda^m Z^m}{m!}$ .

Taking expectation and using  $\mathbb{E}Z = 0$  gives

$$\mathbb{E}e^{\lambda Z} = 1 + \sum_{m=2}^{\infty} \frac{\lambda^m \mathbb{E}[Z^m]}{m!} \leq 1 + \sum_{m=2}^{\infty} \frac{|\lambda|^m \mathbb{E}[|Z|^m]}{m!}, \tag{16}$$

using triangle inequality. Assuming the  $\psi_1$ -integrability, i.e., using (15),

$$\mathbb{E} \exp\left(\frac{|Z|}{K}\right) = \sum_{m=0}^{\infty} \frac{\mathbb{E}[|Z|^m]}{K^m m!} \leq 2,$$

where all the terms in the sum are non-negative. So, for each  $m$ ,

$$\frac{\mathbb{E}[|Z|^m]}{K^m m!} \leq 2 \implies \mathbb{E}[|Z|^m] \leq 2K^m m!.$$

Putting this back in (16) gives:

$$\mathbb{E}e^{\lambda Z} \leq 1 + \sum_{m=2}^{\infty} \frac{|\lambda|^m}{m!} \cdot 2K^m m! = 1 + 2 \sum_{m=2}^{\infty} (|\lambda|K)^m, \tag{17}$$

where the latter geometric sum converges if  $|\lambda|K < 1$ . Now, for  $|\lambda| \leq \frac{1}{2K}$ ,  $1 - 2|\lambda|K \geq 1/2$ , so (17) becomes,

$$\mathbb{E}e^{\lambda Z} \leq 1 + \frac{2\lambda^2 K^2}{1 - 2|\lambda|K} \leq 1 + 4\lambda^2 K^2.$$

Now, taking log and using the fact  $\log(1 + u) \leq u$  for  $u \geq 0$ ,

$$\log \mathbb{E}e^{\lambda Z} \leq \log(1 + 4\lambda^2 K^2) \leq 4\lambda^2 K^2.$$

Hence, there exists absolute constants  $c, C > 0$  such that for all  $|\lambda| \leq \frac{c}{K}$ ,

$$\psi_Z(\lambda) = \log \mathbb{E}e^{\lambda Z} \leq C\lambda^2 K^2. \quad \square \quad (18)$$

(a4) Assume (iv). Show there exists an absolute constant  $c > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P}\{|Z| \geq t\} \leq 2 \exp\left(-c \min\left(\frac{t^2}{K^2}, \frac{t}{K}\right)\right).$$

Deduce in particular (i) (after adjusting constants).

**Solution:**

Using Chernoff bound for the upper tail,

$$\mathbb{P}\{Z \geq t\} \leq \inf_{\lambda > 0} \exp(-\lambda t + \psi_Z(\lambda)) \leq \exp(-\lambda t + C_0 \lambda^2 K^2),$$

for all  $\lambda \in [0, c_0/K]$ , where the latter inequality is true assuming (iv), i.e., local quadratic log-MGF. Let  $f(\lambda) = -\lambda t + C_0 \lambda^2 K^2 \implies f'(\lambda) = -t + 2C_0 \lambda K^2$ . Setting  $f' = 0$  gives

$$-t + 2C_0 \lambda K^2 = 0 \implies \lambda^* = \frac{t}{2C_0 K^2}.$$

Now, we have two possible cases:

- $\lambda^* \leq c_0/K$ , which happens iff  $t \leq 2c_0 C_0 K$ . So, for  $t \leq 2c_0 C_0 K$  the unconstrained optimizer is admissible, and plugging this  $\lambda^*$  into  $f(\lambda)$  gives:

$$f(\lambda^*) = -\frac{t^2}{2C_0 K^2} + C_0 K^2 \cdot \frac{t^2}{4C_0^2 K^4} = -\frac{t^2}{2C_0 K^2} + \frac{t^2}{4C_0 K^2} = -\frac{t^2}{4C_0 K^2}.$$

So, for  $t \lesssim K$  we get Gaussian-type tail, i.e.,

$$\mathbb{P}\{Z \geq t\} \leq \exp\left(-\frac{t^2}{4C_0 K^2}\right). \quad (19)$$

- $\lambda^* > c_0/K$ , i.e. the unconstrained optimizer is outside the feasible region  $[0, c_0/K]$ , which happens iff  $t > 2c_0 C_0 K$ . Here, we use the endpoint,  $\lambda^{**} = c_0/K$ , which gives

$$f(\lambda^{**}) = -\frac{c_0 t}{K} + c_0^2 C_0 \leq -\frac{c_0 t}{K} + \frac{c_0 t}{2K} = -\frac{c_0}{2} \cdot \frac{t}{K},$$

where the inequality is true because for  $t > 2c_0 C_0 K \implies \frac{c_0 t}{2K} \geq c_0^2 C_0$ . So, for  $t \gtrsim K$  we get Exponential-type tail, i.e.,

$$\mathbb{P}\{Z \geq t\} \leq \exp\left(-\frac{c_0}{2} \cdot \frac{t}{K}\right). \quad (20)$$

Combining (19) and (20), there exists an absolute constant  $c > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P}\{Z \geq t\} \leq \exp\left(-c \min\left(\frac{t^2}{k^2}, \frac{t}{K}\right)\right). \quad (21)$$

Now, we apply the same argument to  $-Z$ . Since  $\psi_{-Z}(\lambda) = \psi_Z(-\lambda)$ , assumption (iv) also holds for  $-Z$ , so

$$\mathbb{P}\{-Z \geq t\} \leq \exp\left(-c \min\left(\frac{t^2}{k^2}, \frac{t}{K}\right)\right). \quad (22)$$

Finally, using (21) and (22), and union bound gives:

$$\mathbb{P}\{|Z| \geq t\} \leq \mathbb{P}\{Z \geq t\} + \mathbb{P}\{-Z \geq t\} = 2 \exp\left(-c \min\left(\frac{t^2}{k^2}, \frac{t}{K}\right)\right). \quad \square \quad (23)$$

Now, for  $t \geq K$ ,  $\min\left(\frac{t^2}{k^2}, \frac{t}{K}\right) \geq \frac{t}{2K}$ , so (23) gives

$$\mathbb{P}\{|Z| \geq t\} \leq 2 \exp\left(-c \frac{t}{2K}\right). \quad (24)$$

Similarly, when  $0 \leq t \leq K$ , we know  $2 \exp\left(-\frac{t}{\alpha K}\right) \geq 2 \exp\left(-\frac{1}{\alpha}\right)$ , where

$$2 \exp\left(-\frac{1}{\alpha}\right) \geq 1 \iff \alpha \geq \frac{1}{\log 2}.$$

Thus, for  $\alpha \geq \frac{1}{\log 2}$ ,

$$\mathbb{P}\{|Z| \geq t\} \leq 1 \leq 2 \exp\left(-\frac{t}{\alpha K}\right). \quad (25)$$

Let  $C' = \min\left\{\frac{c}{2}, \frac{1}{\alpha}\right\}$ , then combining (24) and (25), for all  $t \geq 0$ ,

$$\mathbb{P}\{|Z| \geq t\} \leq 2 \exp\left(-C' \frac{t}{K}\right). \quad \square \quad (26)$$

- (b) **[Bonus] Comparison with sub-Gaussianity.** Show that sub-Gaussian implies sub-exponential: if  $Z$  is  $\sigma^2$ -sub-Gaussian, then  $\|Z\|_{\psi_1} \leq C\sigma$ . Give an explicit example of a centered random variable that is sub-exponential but not sub-Gaussian (e.g.  $Z = X - \mathbb{E}X$  with  $X \sim \text{Exp}(1)$ ), and justify both claims from tails.

**[Bonus] Solution:**

Assume  $Z$  is  $(\sigma^2)$ -sub-Gaussian, then

$$\mathbb{P}\{|Z| \geq t\} \leq 2 \exp\left(-\frac{t^2}{2\sigma^2}\right), \quad \forall t \geq 0.$$

For any  $t \geq 0$ ,  $\frac{t^2}{2\sigma^2} \geq \frac{t}{2\sigma}$  whenever  $t \geq \sigma$ . So, for all  $t \geq \sigma$ ,  $\mathbb{P}\{|Z| \geq t\} \leq 2 \exp\left(-\frac{t}{2\sigma}\right)$ .

And, for  $0 \leq t \leq \sigma$ ,

$$\mathbb{P}\{|Z| \geq t\} \leq 1 \leq 2 \exp\left(-\frac{t}{2\sigma}\right),$$

because  $e^{-t/(2\sigma)} \geq e^{-1/2} > 1$ . Hence, for all  $t \geq 0$ ,

$$\mathbb{P}\{|Z| \geq t\} \leq 2 \exp\left(-\frac{t}{2\sigma}\right),$$

i.e.,  $Z$  is sub-exponential and  $\|Z\|_{\psi_1} \leq C\sigma$  for some constant  $C$ . Thus,

$$Z \text{ is sub-Gaussian} \implies Z \text{ is sub-exponential.} \quad \square$$

Now, let  $Z = X - 1$  where  $X \sim \text{Exp}(1)$ , then  $\mathbb{E}Z = 0$  (i.e.,  $Z$  is centered). Then, for  $t \geq 0$ ,

$$\mathbb{P}\{Z \geq t\} = \mathbb{P}\{Z \geq 1 + t\} = e^{-(1+t)} = e^{-1}e^{-t} \leq e^{-t}.$$

We know,  $Z \geq -1$ , so for  $t > 1$ ,  $\mathbb{P}\{Z \leq -t\} = 0$ . Therefore, using the union bound gives

$$\mathbb{P}\{|Z| \geq t\} \leq 2e^{-t}, \quad (t \geq 1).$$

After adjusting for constants,

$$\mathbb{P}\{|Z| \geq t\} \leq 2e^{-ct}, \quad (t \geq 0),$$

so  $Z$  has exponential tails, hence  $Z$  is sub-exponential. Now, for  $Z$  to be sub-Gaussian, there need to exist some  $C > 0$  such that

$$\mathbb{P}\{|Z| \geq t\} \leq 2 \exp\left(-\frac{t^2}{2C^2}\right), \quad (t \geq 0). \quad (27)$$

But for  $t \geq 0$ ,

$$\mathbb{P}\{|Z| \geq t\} \geq \mathbb{P}\{Z \geq t\} = e^{-(t+1)}.$$

So, (27) would force  $\exp(-(t+1)) \leq 2 \exp\left(-\frac{t^2}{2C^2}\right)$  for all  $t \geq 0$ . Taking logs:

$$-(t+1) \leq \log 2 - \frac{t^2}{2C^2} \implies t \geq \frac{t^2}{2C^2} - \log 2 - 1,$$

which is not possible for large  $t$  because the right side is quadratic in  $t$ . So, there does not exist any  $C > 0$  such that (27) holds for all  $t \geq 0$ , hence  $Z$  is not sub-Gaussian.  $\square$

### 3 Bounding $\ell_p$ norms of random vectors

Let  $X = (X_1, \dots, X_n) \in \mathbb{R}^n$ . For  $p \in [1, \infty)$ , define

$$\|X\|_p := \left( \sum_{i=1}^n |X_i|^p \right)^{1/p}, \quad \|X\|_\infty := \max_{1 \leq i \leq n} |X_i|.$$

Throughout, you may freely use the standard sub-Gaussian consequences: there exist universal constants  $c, C > 0$  such that for every  $i$  and every  $q \geq 1$ ,

$$(\mathbb{E}|X_i|^q)^{1/q} \leq CK\sqrt{q}, \quad \text{and} \quad \mathbb{P}\{|X_i| \geq t\} \leq 2 \exp\left(-c \frac{t^2}{K^2}\right) \quad (t \geq 0).$$

(a) **Expected  $\ell_p$  norm (finite  $p$ ).** Show that for every  $p \in [1, \infty)$ ,

$$\mathbb{E} \|X\|_p \leq CK \sqrt{p} n^{1/p}. \tag{28}$$

**Solution:**

$x \mapsto x^{1/p}$  for  $p \geq 1$  is concave, so using Jensen's inequality for concave function gives:

$$\mathbb{E} \|X\|_p = \mathbb{E} \left[ \left( \sum_{i=1}^n |X_i|^p \right)^{1/p} \right] \leq \left( \mathbb{E} \left[ \sum_{i=1}^n |X_i|^p \right] \right)^{1/p} = (nC^p K^p p^{p/2})^{1/p} = CK \sqrt{p} n^{1/p},$$

where we use the identical moment bound for each  $X_i$ .  $\square$

(b) **Expected  $\ell_\infty$  norm.** Prove that there exist universal constants  $c, C > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P}\{\|X\|_\infty \geq CK(\sqrt{\log n} + t)\} \leq 2e^{-ct^2}.$$

Deduce that

$$\mathbb{E} \|X\|_\infty \leq CK\sqrt{\log n}.$$

**Solution:**

Because  $\max_i |X_i| \geq u \implies |X_i| \geq u$  for all  $i$ , using union bound gives:

$$\begin{aligned} \mathbb{P}\{\|X\|_\infty \geq CK(\sqrt{\log n} + t)\} &\leq \sum_{i=1}^n \mathbb{P}\{|X_i| \geq CK(\sqrt{\log n} + t)\} \\ &\leq n \cdot 2 \exp\left(-c_1 C^2 (\sqrt{\log n} + t)^2\right) && \text{[tail bound for } X_i\text{]} \\ &\leq 2 \exp\left(-c_1 C^2 (\log n + t^2) + \log n\right) && \text{[(} a + b \text{)}^2 \leq a^2 + b^2\text{]} \\ &= 2 \exp\left((1 - c_1 C^2) \log n\right) \cdot \exp\left(-c_1 C t^2\right). \end{aligned}$$

Here, if we choose  $1 - c_1 C^2 \leq 0 \iff C \geq 1/\sqrt{c_1}$ , then  $\exp\left((1 - c_1 C^2) \log n\right) \leq 1$ . Hence, there exists universal constants  $C \geq 1/\sqrt{c_1} > 0$  and  $C = c_1 C$  such that for all  $t \geq 0$ ,

$$\mathbb{P}\{\|X\|_\infty \geq CK(\sqrt{\log n} + t)\} \leq 2e^{-ct^2}. \quad \square \tag{29}$$

Now, using the tail integral identity  $\mathbb{E}Y = \int_0^\infty \mathbb{P}\{Y \geq s\} ds$  gives:

$$\mathbb{E} \|X\|_\infty = \int_0^\infty \mathbb{P}\{\|X\|_\infty \geq s\} ds = \int_0^{CK\sqrt{\log n}} \mathbb{P}\{\|X\|_\infty \geq s\} ds + \int_{CK\sqrt{\log n}}^\infty \mathbb{P}\{\|X\|_\infty \geq s\} ds,$$

where we can bound the first integral by  $CK\sqrt{\log n}$  because probability is always  $\leq 1$ . So,

$$\mathbb{E} \|X\|_\infty \leq CK\sqrt{\log n} + \int_{CK\sqrt{\log n}}^\infty \mathbb{P}\{\|X\|_\infty \geq s\} ds.$$

Let  $s = CK(\sqrt{\log n} + t) \implies ds = CK dt$ , then using the tail bound in (29):

$$\mathbb{E} \|X\|_\infty \leq CK\sqrt{\log n} + CK \int_{CK\sqrt{\log n}}^\infty 2e^{-ct^2} dt \leq CK\sqrt{\log n} + 2CK \int_0^\infty e^{-ct^2} dt,$$

where the latter integral is just a constant, as it's an integral of a Gaussian kernel. Then,

$$\mathbb{E} \|X\|_\infty \leq CK\sqrt{\log n} + C_1K \leq C'K\sqrt{\log n}.$$

Hence, there exists universal constant  $C > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{E} \|X\|_\infty \leq CK\sqrt{\log n}. \quad (30)$$

(c) **Two regimes of expected  $\ell_p$  norms.** Show that

$$\mathbb{E} \|X\|_p \leq \begin{cases} CK \sqrt{p} n^{1/p}, & 1 \leq p \leq \log n, \\ CK \sqrt{\log n}, & p \geq \log n, \end{cases}$$

up to universal constants.

**Solution:**

From Exercise 1 in HW1, we have for every  $p \geq 1$ ,  $\|X\|_p \leq n^{1/p} \|X\|_\infty$ . So, using (30):

$$\mathbb{E} \|X\|_p \leq n^{1/p} \|X\|_\infty \leq CK\sqrt{\log n} n^{1/p}. \quad (31)$$

Combining (28) and (31) gives:

$$\mathbb{E} \|X\|_p \leq \min \left\{ CK \sqrt{p} n^{1/p}, CK \sqrt{\log n} n^{1/p} \right\}.$$

When  $p \geq \log n$ ,  $n^{1/p} \leq e$  so we can merge  $n^{1/p}$  into  $C$  in (31) for  $p \geq \log n$ . Hence,

$$\mathbb{E} \|X\|_p \leq \begin{cases} CK \sqrt{p} n^{1/p}, & 1 \leq p \leq \log n, \\ CK \sqrt{\log n}, & p \geq \log n, \end{cases}$$

up to universal constants.  $\square$

- (d) **[Bonus] High-probability bound for  $\ell_p$  norm when  $p \geq \log n$ .** Fix  $p \in [\log n, \infty]$ . Show that there exists universal constants  $c, C > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P} \left\{ \|X\|_p \geq CK \left( \sqrt{\log n} + t \right) \right\} \leq 2e^{-ct^2}.$$

**[Bonus] Solution:**

Similar to part (c), when  $p \geq \log n$ ,  $n^{1/p} \leq e$  so we get  $\|X\|_p \leq e \|X\|_\infty$ . Fix  $p \in [\log n, \infty]$ . Using (29), there exists a universal constants  $c, C_0 > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P} \left\{ \|X\|_\infty \geq C_0 K \left( \sqrt{\log n} + t \right) \right\} \leq 2e^{-ct^2}. \tag{32}$$

Choose  $C = eC_0$ , then

$$\|X\|_p \geq CK \left( \sqrt{\log n} + t \right) \implies \|X\|_\infty \geq \frac{\|X\|_p}{e} \geq C_0 K \left( \sqrt{\log n} + t \right).$$

This gives the event  $A := \left\{ \|X\|_\infty \geq CK \left( \sqrt{\log n} + t \right) \right\} \subseteq \left\{ \|X\|_\infty \geq C_0 K \left( \sqrt{\log n} + t \right) \right\} =: B$ , i.e.  $\mathbb{P} \{A\} \leq \mathbb{P} \{B\}$ . Using (32), for  $p \in [\log n, \infty]$ , there exists  $c, C > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P} \left\{ \|X\|_p \geq CK \left( \sqrt{\log n} + t \right) \right\} \leq 2e^{-ct^2}. \quad \square \tag{33}$$

- (e) **[Bonus] High-probability bound for  $\ell_p$  norm when  $p \leq \log n$ .** Fix  $p \in [1, \log n]$ . Show that for every  $q \geq p$ ,

$$\left( \mathbb{E} \|X\|_p^q \right)^{1/q} \leq CK n^{1/p} \sqrt{q}.$$

Deduce that there exist universal constants  $c, C > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P} \left\{ \|X\|_p \geq CK n^{1/p} \left( \sqrt{p} + t \right) \right\} \leq 2e^{-ct^2}.$$

**[Bonus] Solution:**

Fix  $p \in [1, \log n]$  and define  $Y := \|X\|_p$ . Let  $r = q/p \geq 1$ , then using the identity  $\left( \sum a_i \right)^r \leq n^{r-1} \sum a_i^r$  gives  $Y^q \leq n^{q/p-1} \sum_{i=1}^n |X_i|^q$ . So,

$$\mathbb{E} Y^q \leq n^{q/p-1} \sum_{i=1}^n \mathbb{E} |X_i|^q \leq n^{q/p-1} n (C_1 K \sqrt{q})^q = \left( C_1 K n^{1/p} \sqrt{q} \right)^q,$$

where we use sub-Gaussian moment bound for each  $X_i$ . Hence, for every  $q \geq p$ ,

$$\left( \mathbb{E} \|X\|_p^q \right)^{1/q} \leq C_1 K n^{1/p} \sqrt{q}. \quad \square \tag{34}$$

Let  $q = p + t^2$ , then using Markov's inequality:

$$\mathbb{P} \left\{ Y \geq e \left( \mathbb{E} Y^q \right)^{1/q} \right\} = \mathbb{P} \left\{ Y^q \geq e^q \mathbb{E} [Y^q] \right\} \leq e^{-q} = e^{-(p+t^2)} \leq e^{-t^2}. \tag{35}$$

Using the bound for  $\left( \mathbb{E} Y^q \right)^{1/q}$  from (34) in (35) gives,

$$\mathbb{P} \left\{ Y \geq e C_1 K n^{1/p} \sqrt{p+t^2} \right\} \leq e^{-t^2} \implies \mathbb{P} \left\{ Y \geq e C_1 K n^{1/p} \left( \sqrt{p} + t \right) \right\} \leq e^{-t^2},$$

using the fact that  $\sqrt{p+t^2} \leq \sqrt{p} + t$ . Hence, there exists universal constants  $c, C > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P} \{ \|X\|_p \geq CK n^{1/p} (\sqrt{p} + t) \} \leq 2e^{-ct^2}. \quad \square \tag{36}$$

(f) **[Bonus] Thin shell for  $\|X\|_2$  (sharper than part (d)).** Assume in addition that  $\mathbb{E}X_i^2 = 1$  for all  $i$ . Show that there exist universal constants  $c, C > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P} \{ \left| \|X\|_2 - \sqrt{n} \right| \geq t \} \leq 2 \exp \left( -c \frac{t^2}{K^4} \right).$$

**[Bonus] Solution:**

Here, each  $X_i$  is sub-Gaussian with  $\|X_i\|_{\psi_2} \leq K$ . So,  $X_i^2$  will be sub-exponential with  $\|X_i^2\|_{\psi_1} \leq C_1 \|X_i\|_{\psi_2}^2 \leq C_1 K^2$ , for some constant  $C_1 > 0$ . Let  $Y_i = X_i^2 - 1$ , then  $Y_i$  will also be sub-exponential with  $\|Y_i\|_{\psi_1} \leq C_1 K^2$  and  $Y_i$  will be centered as  $\mathbb{E}X_i^2 = 1$  (from assumption).

Define  $Z := n^{-1} \|X\|_2^2 - 1 = n^{-1} \sum_{i=1}^n X_i^2 - 1 = n^{-1} \sum_{i=1}^n Y_i$ . Here,  $Y_i$ 's are independent, so using Bernstein's inequality for the sum of independent sub-exponential gives,

$$\mathbb{P} \{ |Z| \geq u \} \leq 2 \exp \left\{ -c_1 n \cdot \min \left( \frac{u^2}{K^4}, \frac{u}{K^2} \right) \right\}, \tag{37}$$

for all  $u \geq 0$ , where  $c_1 > 0$  is an absolute constant. Let  $\delta = \frac{t}{\sqrt{n}}$ , then using the implication

$$\left| \frac{\|X\|_2}{\sqrt{n}} - 1 \right| \geq \delta \implies \left| \frac{\|X\|_2^2}{n} - 1 \right| \geq \max \{ \delta, \delta^2 \} \text{ gives,}$$

$$\mathbb{P} \{ \left| \|X\|_2 - \sqrt{n} \right| \geq t \} \leq \mathbb{P} \left\{ |Z| \geq \max \left( \frac{t}{\sqrt{n}}, \frac{t^2}{n} \right) \right\}. \tag{38}$$

Putting  $u = \max \left( \frac{t}{\sqrt{n}}, \frac{t^2}{n} \right)$  in (37), we get

$$\mathbb{P} \{ \left| \|X\|_2 - \sqrt{n} \right| \geq t \} \leq 2 \exp \left\{ -c_1 n \cdot \min \left( \frac{u^2}{K^4}, \frac{u}{K^2} \right) \right\}. \tag{39}$$

Now, we have two cases:

i. When  $t \leq \sqrt{n}$ ,  $u = t/\sqrt{n}$ . Then,

$$c_1 n \cdot \frac{u^2}{K^4} = c_1 \cdot \frac{t^2}{K^4}.$$

Similarly,

$$c_1 n \cdot \frac{u}{K^2} = c_1 \frac{t\sqrt{n}}{K^2} \geq c_1 \cdot \frac{t^2}{K^2} \geq c_1 \cdot \frac{t^2}{K^4}.$$

So, in this case:

$$c_1 n \cdot \min \left( \frac{u^2}{K^4}, \frac{u}{K^2} \right) \geq c_1 \cdot \frac{t^2}{K^4}.$$

ii. When  $t \geq \sqrt{n}$ ,  $u = t^2/n$ . Then,

$$c_1 n \cdot \frac{u^2}{K^4} = c_1 \cdot \frac{t^4}{nK^4} \geq c_1 \cdot \frac{t^2}{K^4},$$

because  $t^2 \geq n$ . Similarly,

$$c_1 n \cdot \frac{u}{K^2} = c_1 \cdot \frac{t^2}{K^2} \geq c_1 \cdot \frac{t^2}{K^4}.$$

So, in this case:

$$c_1 n \cdot \min\left(\frac{u^2}{K^4}, \frac{u}{K^2}\right) \geq c_1 \cdot \frac{t^2}{K^4}.$$

Hence, combining both cases and putting it back in (39), there exists universal constant  $c > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P}\left\{|\|X\|_2 - \sqrt{n}| \geq t\right\} \leq 2 \exp\left(-c \frac{t^2}{K^4}\right). \quad \square$$

## 4 Bounding $L_p$ norms of linear forms (Khintchine inequalities)

Throughout, let  $X_1, \dots, X_n$  be independent random variables and let  $a = (a_1, \dots, a_N) \in \mathbb{R}^n$ . Write

$$S(a) := \sum_{i=1}^N a_i X_i, \quad \|a\|_2 := \left( \sum_{i=1}^N a_i^2 \right)^{1/2}, \quad \|a\|_\infty := \max_{1 \leq i \leq N} |a_i|.$$

Throughout,  $C, c > 0$  denote absolute constants (that may change from line to line).

(a) **Sub-Gaussian Khintchine  $L_p$  bound** ( $p \geq 2$ ). Assume

$$\mathbb{E}X_i = 0, \quad \mathbb{E}X_i^2 = 1, \quad \|X_i\|_{\psi_2} \leq K_2 \quad \text{for all } i,$$

for some  $K_2 \geq 1$ . Show that for every  $p \in [2, \infty)$ ,

$$\|a\|_2 \leq \|S(a)\|_{L_p} \leq CK_2 \sqrt{p} \|a\|_2. \quad (40)$$

### Solution:

Using the monotonicity of  $L_p$ -norm, for  $0 < p \leq q \leq \infty$ ,  $\|X\|_{L_p} \leq \|X\|_{L_q}$ . So, for all  $p \in [2, \infty)$ ,

$$\|S(a)\|_{L_2} \leq \|S(a)\|_{L_p},$$

where

$$\|S(a)\|_{L_2}^2 = \mathbb{E} \left( \sum_{i=1}^N a_i X_i \right)^2 = \sum_{i=1}^N a_i^2 \mathbb{E}[X_i^2] + \sum_{1 \leq i < j \leq N} a_i a_j \mathbb{E}[X_i X_j] = \sum_{i=1}^N a_i^2 = \|a\|_2^2,$$

because  $X_i$ 's are independent,  $\mathbb{E}[X_i X_j] = 0$  for all  $i \neq j$ . So, for all  $p \in [2, \infty)$ ,  $\|a\|_2 \leq \|S(a)\|_{L_p}$ . Because  $X_i$  is Gaussian with  $\|X_i\|_{\psi_2} \leq K_2$ ,  $a_i X_i$  will also be Gaussian with  $\|a_i X_i\|_{\psi_2} \leq |a_i| K_2$ . Now, using the sub-Gaussian Hoeffding inequality for sum of independent Gaussians:

$$\|S(a)\|_{\psi_2}^2 \leq C_1 \sum_{i=1}^N \|a_i X_i\|_{\psi_2}^2 \leq C_1 K_2^2 \sum_{i=1}^N a_i^2 = C_1 K_2^2 \|a\|_2^2 \implies \|S(a)\|_{\psi_2} \leq C_2 K_2 \|a\|_2.$$

$S(a)$  is sub-Gaussian because it's a sum of sub-Gaussians, and for a sub-Gaussian  $Y$ , the relation between moment growth and exponential integrability gives  $\|Y\|_{L_p} \leq C_3 \sqrt{p} \|Y\|_{\psi_2}$ . So,  $\|S(a)\|_{L_p} \leq C \sqrt{p} K_2 \|a\|_2$ , and hence for every  $p \in [2, \infty)$ ,

$$\|a\|_2 \leq \|S(a)\|_{L_p} \leq CK_2 \sqrt{p} \|a\|_2. \quad \square$$

(b) **Sub-exponential Bernstein-Khintchine tail bound.** Assume instead that

$$\mathbb{E}X_i = 0, \quad \|X_i\|_{\psi_1} \leq K_1 \quad \text{for all } i,$$

for some  $K_1 > 0$ . Show that there exists an absolute constant  $c > 0$  such that for all  $t \geq 0$ ,

$$\mathbb{P}\{|S(a)| \geq t\} \leq 2 \exp \left[ -c \min \left( \frac{t^2}{K_1^2 \|a\|_2^2}, \frac{t}{K_1 \|a\|_\infty} \right) \right]. \quad (41)$$

**Solution:**

$X_i$ 's are sub-Exponential with  $\|X_i\|_{\psi_1} \leq K_1$ , so  $\|a_i X_i\|_{\psi_1} \leq |a_i| K_1$ . And,

$$V = \sum_{i=1}^N \|a_i X_i\|_{\psi_1}^2 \leq K_1^2 \sum_{i=1}^N a_i^2 = K_1^2 \|a\|_2^2, \quad \text{and} \quad B = \max_{1 \leq i \leq N} \|a_i X_i\|_{\psi_1} \leq K_1 \max_{1 \leq i \leq N} a_i = K_1 \|a\|_\infty.$$

Using the Bernstein inequality for sub-exponential sums on  $S(a)$ ,  $\exists c > 0$  s.t. for all  $t \geq 0$ ,

$$\mathbb{P}\{|S(a)| \geq t\} \leq 2 \exp\left[-c \min\left(\frac{t^2}{V}, \frac{t}{B}\right)\right] \leq 2 \exp\left[-c \min\left(\frac{t^2}{K_1^2 \|a\|_2^2}, \frac{t}{K_1 \|a\|_\infty}\right)\right]. \quad \square$$

- (c) **Sub-exponential Bernstein-Khintchine  $L_p$  bound** ( $p \geq 2$ ). Use (41) to show that there exists an absolute constant  $C > 0$  such that for every  $p \geq 2$ ,

$$\|S(a)\|_{L_p} \leq CK_1 (\sqrt{p} \|a\|_2 + p \|a\|_\infty). \quad (42)$$

**Solution:**

Let  $\alpha := K_1^2 \|a\|_2^2$  and  $\beta := K_1 \|a\|_\infty$ . Using the tail integral identity and bound in (41) gives

$$\mathbb{E}[|S(a)|^p] \leq p \int_0^\infty t^{p-1} \cdot 2 \exp\left[-c \min\left(\frac{t^2}{\alpha}, \frac{t}{\beta}\right)\right] dt.$$

Using  $e^{-\min\{u,v\}} \leq e^{-u} + e^{-v}$ , we split the integral:

$$\frac{\mathbb{E}[|S(a)|^p]}{2pe^c} \leq \int_0^\infty t^{p-1} e^{-t^2/\alpha} dt + \int_0^\infty t^{p-1} e^{-t/\beta} dt = \frac{1}{2} \int_0^\infty z^{p/2-1} e^{-z/\alpha} dz + \int_0^\infty t^{p-1} e^{-t/\beta} dt,$$

where we do  $z = t^2$  substitution in the first integral. Now, we recognize both integrals as the integrals of a Gamma kernel, which gives

$$\frac{\mathbb{E}[|S(a)|^p]}{2pe^c} \leq \frac{1}{2} \alpha^{p/2} \Gamma(p/2) + \beta^p \Gamma(p).$$

Rearranging the terms:

$$\mathbb{E}[|S(a)|^p] \leq pe^c K_1^p (\|a\|_2^p \Gamma(p/2) + 2 \|a\|_\infty^p \Gamma(p)).$$

Using the bound  $\Gamma(s) \leq s^s$  for  $s \geq 1$  gives  $\Gamma(p/2) \leq (\sqrt{p})^p$  and  $\Gamma(p) \leq p^p$ . So,

$$(\mathbb{E}[|S(a)|^p])^{1/p} \leq C_1 p^{1/p} K_1 (\|a\|_2^p (\sqrt{p})^p + \|a\|_\infty^p p^p)^{1/p} \leq CK_1 (\|a\|_2^p (\sqrt{p})^p + \|a\|_\infty^p p^p)^{1/p},$$

for some absolute constant  $C > 0$  because  $p^{1/p} \leq e^{1/e}$ . Finally, using Minkowski's inequality, there exists an absolute constant  $C > 0$  such that for every  $p \geq 2$ ,

$$\|S(a)\|_{L_p} = (\mathbb{E}[|S(a)|^p])^{1/p} \leq CK_1 (\sqrt{p} \|a\|_2 + p \|a\|_\infty). \quad \square$$

- (d) **[Bonus] The  $p \|a\|_\infty$  term is unavoidable.** Give an example showing that the bound (42) cannot, in general, be improved to  $\|S(a)\|_{L_p} \lesssim K_1 \sqrt{p} \|a\|_2$  for all  $p$ .

**[Bonus] Solution:**

Let  $N = 1, a_1 = 1$  and  $X_1$  be centered Laplace (double-exponential) distribution with scale parameter 1, then here  $S(a) = X_1$ . The density of  $X_1$  is given by:

$$f(x) = \frac{1}{2} e^{-|x|}, \quad x \in \mathbb{R}.$$

We know  $X_1$  is sub-Exponential so  $\|X_1\|_{\psi_1} \asymp 1$ . Now, for  $p \geq 1$ ,

$$\mathbb{E}[|X_1|^p] = 2 \int_0^\infty x^p \cdot \frac{1}{2} e^{-x} dx = \int_0^\infty x^p e^{-x} dx = \Gamma(p+1),$$

recognizing the integral of a Gamma kernel. Using Stirling's approximation,  $\Gamma(p+1) = p! \sim \sqrt{2\pi p} \left(\frac{p}{e}\right)^p$  so  $(\Gamma(p+1))^{1/p} \sim \frac{p}{e}$ . This gives

$$\|S(a)\|_{L_p} = \|X_1\|_{L_p} = (\Gamma(p+1))^{1/p} \asymp p \quad \text{for large } p.$$

Hence, the bound (42) cannot, in general, be improved to

$$\|S(a)\|_{L_p} \lesssim \sqrt{p} \quad \text{for all } p. \quad \square$$

## 5 Bounding quadratic forms (Hanson–Wright inequalities)

Let  $g = (g_1, \dots, g_n) \sim \mathcal{N}(0, I_n)$  have independent standard normal coordinates, and let  $A \in \mathbb{R}^{n \times n}$  be a symmetric matrix with zero diagonal ( $a_{ii} = 0$  for all  $i$ ). Define the second-order Gaussian chaos

$$Z := g^\top A g = \sum_{i,j=1}^n a_{ij} g_i g_j = 2 \sum_{1 \leq i < j \leq n} a_{ij} g_i g_j.$$

Let  $\|\cdot\|_{\text{op}}$  denote the spectral/operator norm and  $\|\cdot\|_F$  the Frobenius norm:

$$\|A\|_{\text{op}} := \sup_{\|x\|_2=1} \|Ax\|_2, \quad \|A\|_F := \left( \sum_{i,j=1}^n a_{ij}^2 \right)^{1/2}.$$

(For symmetric  $A$ ,  $\|A\|_{\text{op}}$  is the largest absolute eigenvalue, and  $\|A\|_F^2$  is the sum of squared eigenvalues.)

(a) **Mean and variance.** Compute  $\mathbb{E}Z$  and  $\text{Var}Z$  explicitly, and show that

$$\mathbb{E}Z = 0, \quad \text{Var}(Z) = 2 \|A\|_F^2.$$

### Solution:

Here,  $\mathbb{E}[g_i g_j] = 0$  for all  $i \neq j$  because  $g_i \perp g_j$  for all  $i \neq j$ , and  $\mathbb{E}[g_i] = 0 \forall i$ . Then,

$$\begin{aligned} \mathbb{E}Z &= \mathbb{E} \left[ \sum_{i,j=1}^n a_{ij} g_i g_j \right] \\ &= \underbrace{\mathbb{E} \left[ \sum_{i=j} a_{ii} g_i g_j \right]}_{=0 \text{ b/c } a_{ii}=0 \forall i} + \underbrace{\mathbb{E} \left[ \sum_{i \neq j} a_{ij} g_i g_j \right]}_{\mathbb{E}[g_i g_j]=0 \forall i \neq j} \\ &= 0. \quad \square \end{aligned}$$

We know,  $\text{Var}(Z) = \mathbb{E}Z^2 - (\mathbb{E}Z)^2 = \mathbb{E}Z^2$  because  $\mathbb{E}Z = 0$  and  $Z^2 = 4 \sum_{i < j} \sum_{k < l} a_{ij} a_{kl} g_i g_j g_k g_l$ . So,

$$\text{Var}(Z) = \mathbb{E} \left[ 4 \sum_{i < j} \sum_{k < l} a_{ij} a_{kl} g_i g_j g_k g_l \right] = 4 \sum_{i < j} \sum_{k < l} a_{ij} a_{kl} \mathbb{E}[g_i g_j g_k g_l].$$

Here,  $\mathbb{E}[g_i g_j g_k g_l] = 0$  for all the  $(i, j) \neq (k, l)$  pairs because it will have at least one unpaired independent Gaussian and the whole expectation will go to 0. For  $(i, j) = (k, l)$  pairs,  $\mathbb{E}[g_i g_j g_k g_l] = \mathbb{E}[g_i^2] \mathbb{E}[g_j^2] = 1$ . And, for a symmetric matrix  $A$  with  $\text{diag}(A) = 0$ ,  $\|A\|_F^2 = 2 \sum_{i < j} a_{ij}^2$ . Hence,

$$\text{Var}(Z) = 4 \sum_{i < j} \sum_{k < l} a_{ij} a_{kl} \mathbb{E}[g_i g_j g_k g_l] = 4 \sum_{i < j} a_{ij}^2 = 2 \left( 2 \sum_{i < j} a_{ij}^2 \right) = 2 \|A\|_F^2. \quad \square$$

(b) **Variance bound using Gaussian Poincaré.** Let  $f(x) = x^\top A x$ . Use the Gaussian Poincaré inequality

$$\text{Var}(f(g)) \leq \mathbb{E} \|\nabla f(g)\|_2^2$$

to prove that  $\text{Var}(Z) \leq 4 \|A\|_F^2$ . Compare this bound with your exact formula from part (a). Where did the extra factor come from?

**Solution:**

For  $f(x) = x^T A x$ ,  $\nabla f(x) = Ax + A^T x = 2Ax$ , because  $A^T = A$  for a symmetric matrix. Then,

$$\|\nabla f(g)\|_2^2 = \|2Ag\|_2^2 = 4(Ag)^T(Ag) = 4g^T A^2 g.$$

Taking the expectation gives,

$$\mathbb{E} \|\nabla f(g)\|_2^2 = 4\mathbb{E}[g^T A^2 g] = 4 \cdot \text{trace}(A^2) = 4 \|A\|_F^2.$$

Since,  $g$  is Gaussian, using Gaussian Poincaré inequality gives,

$$\text{Var}(f(g)) \leq \mathbb{E} \|\nabla f(g)\|_2^2 = 4 \|A\|_F^2. \quad \square$$

Comparing to the exact formula from (a), the Poincaré inequality is off by a factor of 2. For a quadratic  $f$ , the gradient grows linearly in  $g$ , and the square of that gives a factor of 4. However, the true variance depends on the fourth Gaussian moment. So, Poincaré is a first-order inequality while quadratic functions fluctuate according to second-order effects, and thus we get an extra factor of 2. Also, the Gaussian Poincaré inequality is a general-purpose bound for any differentiable function. It doesn't "know" that  $f(x)$  is a quadratic form or that  $A$  has a zero diagonal, which ends up giving looser bound.  $\square$

- (c) **Diagonalization and distributional representation.** Let  $\lambda_1, \dots, \lambda_n$  be the eigenvalues of  $A$ . Show that

$$Z \stackrel{d}{=} \sum_{i=1}^n \lambda_i (g_i^2 - 1).$$

**Solution:**

$A$  is a symmetric matrix, so there exists an orthogonal matrix  $U$  such that  $A = U^T \Lambda U$ , where  $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_n)$ . Let  $y = Ug$  then  $\mathbb{E}[y] = 0$  and  $\text{Var}(y) = U^T \text{Var}(g)U = U^T U = I_n$ , which gives rotational invariance, i.e.,  $y \sim \mathcal{N}(0, I_n)$ . Now,

$$Z = g^T A g = g^T U^T \Lambda U g = y^T \Lambda y = \sum_{i=1}^n \lambda_i y_i^2.$$

Here,  $y^T y = g^T U^T U g = g^T g \implies y_i^2 \stackrel{d}{=} g_i^2 \forall i$ , and  $\text{trace}(A) = \sum_i \lambda_i = 0$ . Hence,

$$Z \stackrel{d}{=} \sum_{i=1}^n \lambda_i g_i^2 - \sum_{i=1}^n \lambda_i \stackrel{d}{=} \sum_{i=1}^n \lambda_i (g_i^2 - 1). \quad \square$$

- (d) **CGF of a centered  $\chi_1^2$ .** Let  $G \sim \mathcal{N}(0, 1)$  and define  $Y := G^2 - 1$ . Show that for every  $\theta < \frac{1}{2}$ ,

$$\log \mathbb{E} e^{\theta Y} = -\theta - \frac{1}{2} \log(1 - 2\theta).$$

Then prove the bound

$$\log \mathbb{E} e^{\theta Y} \leq \frac{\theta^2}{1 - (2\theta)_+} \quad \text{for all } \theta < \frac{1}{2},$$

where  $(u)_+ := \max\{u, 0\}$ .

**Solution:**

Here,  $\mathbb{E}e^{\theta Y} = \mathbb{E} \exp((G^2 - 1)\theta) = e^{-\theta} \mathbb{E}e^{\theta G^2}$ . Now, using the definition of expectation:

$$\mathbb{E}e^{\theta Y} = e^{-\theta} \int_{-\infty}^{\infty} e^{\theta x^2} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = \frac{e^{-\theta}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}(1-2\theta)} dx,$$

where the integral converges iff  $1 - 2\theta > 0$ , i.e.,  $\theta < 1/2$ . Recognizing the Gaussian integral with variance  $\sigma^2 = (1 - 2\theta)^{-1}$  gives:

$$\mathbb{E}e^{\theta Y} = \frac{e^{-\theta}}{\sqrt{2\pi}} \cdot \sqrt{2\pi} \cdot (1 - 2\theta)^{-1/2} = e^{-\theta} (1 - 2\theta)^{-1/2}. \quad (43)$$

Finally, taking logarithm of (46) gives the exact CGF,

$$\log \mathbb{E}e^{\theta Y} = -\theta - \frac{1}{2} \log(1 - 2\theta), \quad \forall \theta < \frac{1}{2}. \quad \square \quad (44)$$

We use Taylor expansion of  $-\log(1 - u) = \sum_{k=1}^{\infty} \frac{u^k}{k}$ , valid for  $|u| < 1$ . When  $\theta \geq 0$ , let  $u = 2\theta < 1$ :

$$-\frac{1}{2} \log(1 - 2\theta) = \frac{1}{2} \sum_{k=1}^{\infty} \frac{(2\theta)^k}{k} = \theta + \sum_{k=2}^{\infty} \frac{2^{k-1}\theta^k}{k}.$$

Substituting it back into the exact CGF expression in (44):

$$\log \mathbb{E}e^{\theta Y} = -\theta + \left( \theta + \sum_{k=2}^{\infty} \frac{2^{k-1}\theta^k}{k} \right) = \sum_{k=2}^{\infty} \frac{2^{k-1}\theta^k}{k} \leq \sum_{k=2}^{\infty} \frac{2^{k-1}\theta^k}{2} = \theta^2 \sum_{k=0}^{\infty} (2\theta)^k = \frac{\theta^2}{1 - 2\theta},$$

where the last infinite sum is a Geometric series with ratio  $2\theta < 1$ . When  $\theta < 0$ , let  $x = -2\theta > 0$  and substitute  $x = -2\theta$  in equation (44), which gives  $\log \mathbb{E}e^{\theta Y} \leq (x - \log(1 + x))/2$ . We know,

$$\log(1 + x) \geq x - \frac{x^2}{2} \iff x - \log(1 + x) \leq \frac{x^2}{2}, \quad \text{for all } x \geq 0.$$

Therefore,

$$\log \mathbb{E}e^{\theta Y} \leq \frac{1}{2} \cdot \frac{x^2}{2} = \frac{1}{4} \cdot 4\theta^2 = \theta^2, \quad (\theta < 0)$$

Hence, we get the bound:

$$\log \mathbb{E}e^{\theta Y} \leq \frac{\theta^2}{1 - (2\theta)_+}, \quad \forall \theta < \frac{1}{2}. \quad \square \quad (45)$$

- (e) **[Bonus] A Bernstein-type CGF bound.** Using parts (c) and (d) plus independence, show that for all  $\lambda$  satisfying

$$|\lambda| < \frac{1}{2\|A\|_{\text{op}}},$$

the log-MGF of  $Z$  satisfies the ‘‘Bernstein form’’

$$\log \mathbb{E}e^{\lambda Z} \leq \frac{\lambda^2 \|A\|_F^2}{1 - 2|\lambda| \|A\|_{\text{op}}}. \quad (46)$$

**[Bonus] Solution:**

Using the expression for  $Z$  from part (c):

$$\log \mathbb{E} e^{\theta Z} = \log \mathbb{E} \exp \left( \sum_{i=1}^n \lambda_i (g_i^2 - 1) \right) = \sum_{i=1}^n \log \mathbb{E} e^{\lambda \lambda_i (g_i^2 - 1)}, \quad (47)$$

because the  $g_i$ 's are independent, so the expectation of the product will be the product of the expectations, and taking the logarithm results in the sum of log expectations. Here, each  $g_i \sim \mathcal{N}(0, 1)$  and let  $\theta = \lambda \lambda_i$  then  $|\lambda \lambda_i| < 1/2 \implies \theta < 1/2$ . So,  $|\lambda \lambda_i| \leq |\lambda| \|A\|_{\text{op}} < 1/2 \implies |\lambda| < \frac{1}{2 \|A\|_{\text{op}}}$  suffices  $\theta < 1/2$ . Using the bound (45) from part (d) in (47), for all  $|\lambda| < \frac{1}{2 \|A\|_{\text{op}}}$ ,

$$\log \mathbb{E} e^{\theta Z} \leq \sum_{i=1}^n \frac{(\lambda \lambda_i)^2}{1 - (2\lambda \lambda_i)_+}.$$

We know  $(\lambda \lambda_i)_+ \leq |\lambda \lambda_i| \leq |\lambda| \|A\|_{\text{op}}$  and  $\|A\|_{\text{F}}^2 = \sum_{i=1}^n \lambda_i^2$ . Hence, for all  $|\lambda| < \frac{1}{2 \|A\|_{\text{op}}}$ ,

$$\log \mathbb{E} e^{\theta Z} \leq \sum_{i=1}^n \frac{(\lambda \lambda_i)^2}{1 - 2|\lambda| \|A\|_{\text{op}}} = \frac{\lambda^2}{1 - 2|\lambda| \|A\|_{\text{op}}} \sum_{i=1}^n \lambda_i^2 = \frac{\lambda^2 \|A\|_{\text{F}}^2}{1 - 2|\lambda| \|A\|_{\text{op}}}. \quad \square$$

(f) **[Bonus] Tail bound.** Use the Chernoff bound together with (46) to prove that for all  $u \geq 0$ ,

$$\mathbb{P} \{ Z \geq 2 \|A\|_{\text{F}} \sqrt{u} + 2 \|A\|_{\text{op}} u \} \leq e^{-u}.$$

Deduce the more ‘‘Bernstein-looking’’ bound: for all  $t \geq 0$ ,

$$\mathbb{P} \{ Z \geq t \} \leq \exp \left( - \frac{t^2/4}{\|A\|_{\text{F}}^2 + \|A\|_{\text{op}} t} \right).$$

Finally, show how to turn this into a two-sided bound for  $|Z|$ .

**[Bonus] Solution:**

Using Chernoff and (46), noting  $|\lambda| = \lambda$  for  $\lambda \geq 0$ , on upper tail gives

$$\mathbb{P} \{ Z \geq t \} \leq \inf_{\lambda > 0} \exp \left( -\lambda t + \log \mathbb{E} e^{\lambda Z} \right) \leq \inf_{0 < \lambda < \frac{1}{2 \|A\|_{\text{op}}}} \exp \left( -\lambda t + \frac{\lambda^2 \|A\|_{\text{F}}^2}{1 - 2\lambda \|A\|_{\text{op}}} \right).$$

Choose  $\lambda = \frac{\sqrt{u}}{\|A\|_{\text{F}} + 2 \|A\|_{\text{op}} \sqrt{u}}$ , then  $2 \|A\|_{\text{op}} \lambda = \frac{2 \|A\|_{\text{op}} \sqrt{u}}{\|A\|_{\text{F}} + 2 \|A\|_{\text{op}} \sqrt{u}} < 1$ , so it is admissible.

Then,

$$\begin{aligned} -\lambda t + \frac{\lambda^2 \|A\|_{\text{F}}^2}{1 - 2\lambda \|A\|_{\text{op}}} &= \frac{-t\sqrt{u}}{\|A\|_{\text{F}} + 2 \|A\|_{\text{op}} \sqrt{u}} + \frac{u}{(\|A\|_{\text{F}} + 2 \|A\|_{\text{op}} \sqrt{u})^2} \cdot \frac{\|A\|_{\text{F}}^2 (\|A\|_{\text{F}} + 2 \|A\|_{\text{op}} \sqrt{u})}{\|A\|_{\text{F}}} \\ &= \frac{-t\sqrt{u}}{\|A\|_{\text{F}} + 2 \|A\|_{\text{op}} \sqrt{u}} + \frac{u \|A\|_{\text{F}}}{\|A\|_{\text{F}} + 2 \|A\|_{\text{op}} \sqrt{u}} = \frac{-t\sqrt{u} + u \|A\|_{\text{F}}}{\|A\|_{\text{F}} + 2 \|A\|_{\text{op}} \sqrt{u}}. \end{aligned}$$

Substituting  $t = 2 \|A\|_F \sqrt{u} + 2 \|A\|_{\text{op}} u = \sqrt{u}(2 \|A\|_F + 2 \|A\|_{\text{op}} \sqrt{u})$  gives,

$$-\lambda t + \frac{\lambda^2 \|A\|_F^2}{1 - 2\lambda \|A\|_{\text{op}}} = -\frac{u(2 \|A\|_F + 2 \|A\|_{\text{op}} \sqrt{u} - \|A\|_F)}{\|A\|_F + 2 \|A\|_{\text{op}} \sqrt{u}} = -u.$$

Hence, for all  $u \geq 0$ ,

$$\mathbb{P}\{Z \geq 2 \|A\|_F \sqrt{u} + 2 \|A\|_{\text{op}} u\} \leq e^{-u}. \quad \square \quad (48)$$

Now, let  $u = \frac{t^2/4}{\|A\|_F^2 + \|A\|_{\text{op}} t} \implies t^2 - 4 \|A\|_{\text{op}} u t - 4 \|A\|_F^2 u = 0$ . Solving for  $t \geq 0$  gives,

$$\begin{aligned} t &= \frac{4 \|A\|_{\text{op}} u + \sqrt{16 \|A\|_{\text{op}}^2 u^2 + 16 \|A\|_F^2 u}}{2} = 2 \|A\|_{\text{op}} u + 2\sqrt{\|A\|_{\text{op}}^2 u^2 + \|A\|_F^2 u} \\ &\geq 2 \|A\|_{\text{op}} u + 2\sqrt{\|A\|_F^2 u} \quad [\because \|A\|_{\text{op}}^2 u^2 \geq 0] \\ &= 2 \|A\|_{\text{op}} u + 2 \|A\|_F \sqrt{u}. \end{aligned}$$

Thus, the event  $\{Z \geq t\} \subseteq \{Z \geq 2 \|A\|_{\text{op}} u + 2 \|A\|_F \sqrt{u}\}$ . Hence, using (48), for all  $t \geq 0$ ,

$$\mathbb{P}\{Z \geq t\} \leq \mathbb{P}\{Z \geq 2 \|A\|_F \sqrt{u} + 2 \|A\|_{\text{op}} u\} \leq e^{-u} = \exp\left(-\frac{t^2/4}{\|A\|_F^2 + \|A\|_{\text{op}} t}\right). \quad \square \quad (49)$$

Now, we apply the same argument for  $-Z = g^\top(-A)g$ . We know,  $\|-A\|_F = \|A\|_F$  and  $\|-A\|_{\text{op}} = \|A\|_{\text{op}}$ , the same one-sided bound holds, i.e.,

$$\mathbb{P}\{Z \leq -t\} = \mathbb{P}\{Z \geq -t\} \leq \exp\left(-\frac{t^2/4}{\|A\|_F^2 + \|A\|_{\text{op}} t}\right).$$

Finally, using union bound we get, for all  $t \geq 0$ ,

$$\mathbb{P}\{|Z| \geq t\} \leq 2 \exp\left(-\frac{t^2/4}{\|A\|_F^2 + \|A\|_{\text{op}} t}\right). \quad \square \quad (50)$$

(g) **[Bonus]  $L_p$  bound.** Show that there is an absolute constant  $C$  such that for every integer  $p \geq 2$ ,

$$\|Z\|_{L_p} \leq C(\sqrt{p} \|A\|_F + p \|A\|_{\text{op}}).$$

**[Bonus] Solution:**

Observe, for any  $t \geq 0$ ,

$$\|A\|_F^2 + \|A\|_{\text{op}} t \leq 2 \max(\|A\|_F^2, \|A\|_{\text{op}} t) \implies -\frac{t^2/4}{\|A\|_F^2 + \|A\|_{\text{op}} t} \leq -\frac{1}{8} \min\left(\frac{t^2}{\|A\|_F^2}, \frac{t}{\|A\|_{\text{op}}}\right).$$

So, the tail bound in (50) becomes

$$\mathbb{P}\{|Z| \geq t\} \leq 2 \exp\left(-\frac{1}{8} \min\left(\frac{t^2}{\|A\|_F^2}, \frac{t}{\|A\|_{\text{op}}}\right)\right). \quad (51)$$

Now, using the tail integral identity,

$$\|Z\|_{L_p}^p = \mathbb{E}|Z|^p = p \int_0^\infty t^{p-1} \mathbb{P}\{|Z| \geq t\} dt.$$

Using the bound in (51),

$$\|Z\|_{L_p}^p \leq 2p \int_0^\infty t^{p-1} \exp\left(-\frac{1}{8} \min\left(\frac{t^2}{\|A\|_F^2}, \frac{t}{\|A\|_{\text{op}}}\right)\right) dt.$$

Using  $e^{-\min\{u,v\}} \leq e^{-u} + e^{-v}$ , we split the integral:

$$\|Z\|_{L_p}^p \leq 2p \int_0^\infty t^{p-1} \exp\left(-\frac{t^2}{8\|A\|_F^2}\right) dt + 2p \int_0^\infty t^{p-1} \exp\left(-\frac{t}{8\|A\|_{\text{op}}}\right) dt. \quad (52)$$

Using the Gamma-type estimates (for  $\alpha > 0$ ):

$$\int_0^\infty t^{p-1} e^{-\alpha t^2} dt = \frac{1}{2} \alpha^{-p/2} \Gamma\left(\frac{p}{2}\right), \quad \int_0^\infty t^{p-1} e^{-\alpha t} dt = \alpha^{-p} \Gamma(p),$$

in (52), we get,

$$\|Z\|_{L_p}^p \leq 2p \cdot \frac{1}{2} (8\|A\|_F^2)^{p/2} \Gamma\left(\frac{p}{2}\right) + 2p \cdot (8\|A\|_{\text{op}})^p \Gamma(p) = p8^{p/2} \|A\|_F^p \Gamma\left(\frac{p}{2}\right) + 2p(8\|A\|_{\text{op}})^p \Gamma(p).$$

For  $p \geq 2$ , using the crude bounds  $\Gamma(p/2) \leq (p/2)^{p/2}$  and  $\Gamma(p) \leq p^p$ ,

$$\|Z\|_{L_p}^p \leq p8^{p/2} \|A\|_F^p \left(\frac{p}{2}\right)^{p/2} + 2p(8\|A\|_{\text{op}})^p p^p.$$

Using Minkowski's inequality,

$$\|Z\|_{L_p} \leq \left[ p8^{p/2} \|A\|_F^p \left(\frac{p}{2}\right)^{p/2} \right]^{1/p} + \left[ 2p(8\|A\|_{\text{op}})^p p^p \right]^{1/p} \leq C_1 p^{1/p} \|A\|_F \sqrt{p} + C_2 p^{1/p} \|A\|_{\text{op}} p,$$

for some constants  $C_1, C_2 > 0$ . We know,  $p^{1/p} \leq e^{1/e}$ , so

$$\|Z\|_{L_p} \leq C'_1 \|A\|_F \sqrt{p} + C'_2 \|A\|_{\text{op}} p \leq \max\{C'_1, C'_2\} (\sqrt{p} \|A\|_F + p \|A\|_{\text{op}}).$$

Hence, there is an absolute constant  $C$  such that for every integer  $p \geq 2$ ,

$$\|Z\|_{L_p} \leq C (\sqrt{p} \|A\|_F + p \|A\|_{\text{op}}). \quad \square$$

\*\*\* END OF SOLUTIONS \*\*\*